# On the 'triples test' for symmetry 

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#### Abstract

Randles et al. (1980) and Davis and Quade (1978) independently proposed an asymptotically distribution-free test for testing symmetry about an unknown point. In this note, it has been shown that if $X$ is convex ordered with respect to $-X$, then the one-sided test is consistent, thus solving an open problem of Randles et al. (1980).


Keywords: Symmetry, U-statistic, convex-ordering, nonparametric test.

## 1. Introduction

Let $X_{1}, \ldots, X_{n}$ be a random sample of size $n$ from a continuous distribution with cumulative distribution function $F$ and density function $f$. A random variable $X$ is said to be symmetric about $\theta$ if and only if $X-\theta$ and $\theta-X$ have the same distribution.

Randles, Fligner, Policello and Wolfe (1980) and Davis and Quade (1978) independently proposed an asymptotically distribution-free test for testing the symmetry of a random variable about an unknown point. Their test, which is commonly known as the 'triples test', is based on the U-statistic estimator $\hat{\eta}$ of the parameter

$$
\begin{align*}
\eta= & \operatorname{Pr}\left[X_{1}+X_{2}-2 X_{3}>0\right] \\
& -\operatorname{Pr}\left[X_{1}+X_{2}-2 X_{3}<0\right] \tag{1.1}
\end{align*}
$$

where $X_{1}, X_{2}$ and $X_{3}$ are three independent copies of $X$. Under the null hypothesis of symmetry (about an unknown point), $\eta=0$.

The above authors proposed the 'triples test' purely on heuristic grounds. Randles et al. (1980)

[^0]comment on page 169, "unfortunately, there is no apparent way to express simply the consistency condition that $\eta$ be different from zero in terms of $F^{\prime \prime}$. We attempt to answer this question in the next section and give a general class of distributions for which $\eta$ is different from zero.

## 2. The main result

We shall use the concept of convex ordering of distributions as introduced by van Zwet (1964).

Definition 2.1. Let $X(Y)$ be a random variable with distribution function $F(G)$. Then $X$ is said to be convex ordered with respect to $Y$ if
$G^{-1} F(x)$ is strictly convex on the support of $F$.

We write this as $X \prec^{c} Y$ or $F \prec^{c} G$. This partial ordering of distributions is location-scale invariant in the sense that $F={ }^{\mathrm{c}} G$ if and only if $F(x)$ $=G(\alpha x+\beta)$ for some constants $\alpha$ and $\beta$.

Ahmad and Kochar (1990) proved the following result:

Theorem 2.2. Let $X_{1}^{\prime}, X_{2}^{\prime}$ and $X_{3}^{\prime}\left(Y_{1}^{\prime}, Y_{2}^{\prime}, Y_{3}^{\prime}\right)$ be three independent copies of $X^{\prime}\left(Y^{\prime}\right)$ with distribution function $F(G)$ and let $\lambda$ be any real number such that $0 \leqslant \lambda \leqslant 1$. Then $F \prec^{c} G$ implies

$$
\begin{aligned}
& \operatorname{Pr}\left[X_{1}^{\prime}-\lambda X_{2}^{\prime}-\bar{\lambda} X_{3}^{\prime}<0\right] \\
& \quad<\operatorname{Pr}\left[Y_{1}^{\prime}-\lambda Y_{2}^{\prime}-\bar{\lambda} Y_{3}^{\prime}<0\right],
\end{aligned}
$$

where $\bar{\lambda}=1-\lambda$.
Specializing this result to $X^{\prime}=X-\theta$ and $Y^{\prime}$ $=\theta-X$, we get the following corollary.

Corollary 2.3. Let $X-\theta<^{c} \theta-X$ (or equivalently, $X<^{\mathfrak{c}}-X$ ), then

$$
\begin{aligned}
& \operatorname{Pr}\left[X_{1}-\lambda X_{2}-\bar{\lambda} X_{3}<0\right] \\
& \quad<\operatorname{Pr}\left[X_{1}-\lambda X_{2}-\bar{\lambda} X_{3}>0\right] .
\end{aligned}
$$

In particular, by taking $\lambda=\bar{\lambda}=\frac{1}{2}$, we find that $X<^{\text {c }}-X$ implies $\eta<0$.

Thus the test based on small values of $\hat{\eta}$ or its studentized version will be consistent for testing the null hypothesis

## $\mathrm{H}_{0}: \quad X$ is symmetric

(about an unknown point $\theta$ )
against the one-sided alternative
$\mathrm{H}_{\mathrm{A}}: \quad X-\theta \prec^{\mathrm{c}} \theta-X$
(or equivalently, $X \prec^{\mathrm{c}}-X$ ).
It can be seen that the alternative $\mathrm{H}_{\mathrm{A}}$ is equivalent to
$\frac{f\left[F^{-1}(1-u)\right]}{f\left[F^{-1}(u)\right]}$ nonincreasing in $u$
for $u \in(0,1)$.
The following are some important classes of distributions for which $H_{A}$ holds.

Example 2.4. Let $Y$ be a random variable symmetric about 0 and let, for some $c>0$,
$X_{\theta}= \begin{cases}Y+\theta & \text { if } Y<\theta, \\ Y / c+\theta & \text { if } Y \geqslant \theta .\end{cases}$
Then it can be seen that $X_{\theta}$ satisfies $\mathrm{H}_{\mathrm{A}}$ for any real $\theta$ provided $c<1$.

Example 2.5. Let $g(\cdot)$ be a density function on $[0, \infty$ ). Define a new density function $f$ by
$f(x)= \begin{cases}\frac{1}{2} g(-x) & \text { if } x \leqslant 0, \\ (1 /(2 c)) g(x / c) & \text { if } x>0,\end{cases}$
where $0<c<1$. Then $H_{A}$ holds for such a distribution $f$ for any arbitrary density function $g$ on $[0, \infty)$.

Remarks. (1) It should be noted that $X<^{c}-X$ is not a necessary condition for $\eta<0$. Similarly, it can be shown that if $-X \prec^{c} X$ then $\eta>0$ and the test based on large values of $\hat{\eta}$ will be consistent for testing $\mathrm{H}_{0}$ against the alternative $-X<^{c}$ $X$.
(2) By taking different values of $\lambda$ between 0 and 1 , one can obtain a class of asymptotically distribution-free tests for testing symmetry against $\mathrm{H}_{\mathrm{A}}$, thus generalizing the above test based on $\hat{\eta}$. However, we shall not pursue this matter further here.

## References

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