

## Satterthwaite's Approximate F test

Stat 526  
4-10-25

(1)

Construct 2 new MS terms

$$\begin{aligned} MS' &= MS_t + \dots + MS_u \\ MS'' &= MS_v + \dots + MS_w \end{aligned}$$

no terms in common

So that  $E(MS') - E(MS'') =$  Constant times the desired Variance Component

For yesterday's example

Try  $MS' = MS_A + MS_{ABc}$

$$MS'' = MS_{AB} + MS_{AC}$$

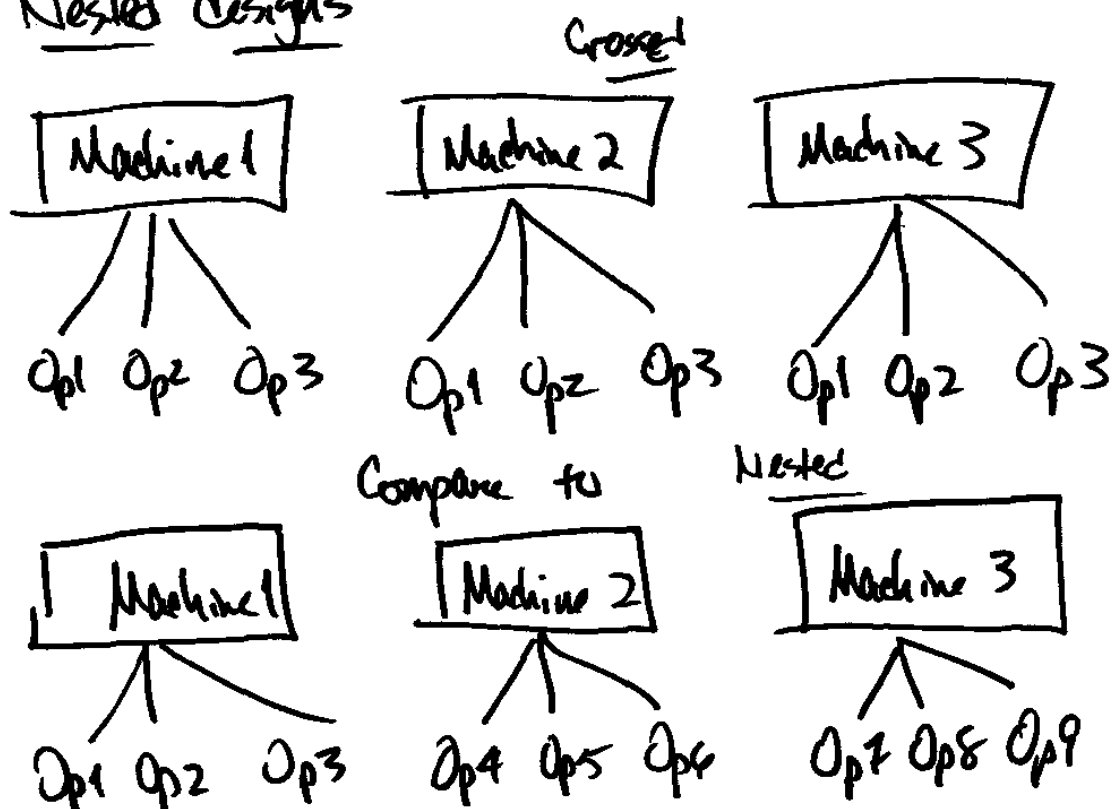
(2)

Satterthwaite proved that  $\frac{MS'}{MS''}$  has

an approximate F distribution with  $p \neq q$  df,

where  $p = \frac{(MS')^2}{\frac{MS_t^2}{df_t} + \dots + \frac{MS_u^2}{df_u}}, q = \frac{(MS'')^2}{\frac{MS_v^2}{df_v} + \dots + \frac{MS_w^2}{df_w}}$

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Nested designs

(4)

Model for a nested design:

$$y_{ijk} = \mu + \underset{\substack{\uparrow \\ \text{machine}}}{\tau_i} + \underset{\substack{\uparrow \\ \text{operator, nested} \\ \text{within machine}}}{\beta_{j(i)}} + \epsilon_{(ij)k}$$

Parameter estimates, assuming both effects are fixed

Assume  $\sum_{i=1}^a \tau_i = 0$  and  $\sum_{j=1}^b \beta_{j(i)} = 0 \quad \forall i$

$$SSE = \sum_i \sum_j \sum_k [y_{ijk} - (\mu + \tau_i + \beta_{j(i)})]^2$$

$$\frac{\partial SSE}{\partial \mu} = \sum_i \sum_j \sum_k 2[y_{ijk} - (\mu + \tau_i + \beta_{j(i)})](-1) \stackrel{\text{set}}{=} 0 \quad (5)$$

$$y_{...} - N\mu - 0 - 0 = 0$$

$$\therefore \hat{\mu} = \frac{y_{...}}{N} = \bar{y}_{...}$$

$$\frac{\partial SSE}{\partial \tau_i} = \sum_j \sum_k 2[y_{ijk} - (\mu + \tau_i + \beta_{j(i)})](-1) \stackrel{\text{set}}{=} 0$$

$$y_{i..} - bn\mu - bn\tau_i - 0 = 0$$

$$\hat{\tau}_i = \frac{y_{i..} - bn\bar{y}_{...}}{bn} = \bar{y}_{i..} - \bar{y}_{...}$$

$$\frac{\partial SSE}{\partial \beta_{j(i)}} = \sum_k 2[y_{ijk} - (\mu + \tau_i + \beta_{j(i)})](-1) \stackrel{\text{set}}{=} 0 \quad (6)$$

$$y_{ij.} - n\mu - n\tau_i - n\beta_{j(i)} = 0$$

$$\hat{\beta}_{j(i)} = \frac{y_{ij.} - n\bar{y}_{...} - n(\bar{y}_{i..} - \bar{y}_{...})}{n}$$

$$= \bar{y}_{ij.} - \bar{y}_{...} - \bar{y}_{i..} + \bar{y}_{...}$$

$$= \bar{y}_{ij.} - \bar{y}_{i..}$$

Compare to the crossed model:

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$$\hat{\mu} = \bar{y}_{...}, \quad \hat{\tau}_i = \bar{y}_{i..} - \bar{y}_{...}$$

$$\hat{\beta}_j = \bar{y}_{.j.} - \bar{y}_{...}, \quad \hat{\tau\beta}_{ij} = \bar{y}_{ij.} - \bar{y}_{i..} - \bar{y}_{.j.} + \bar{y}_{...}$$

SS & df, using rule from Tuesday's class

$$df_A = a - 1$$

$\tau_i$

$$df_{B(A)} = (b-1)a = ab - a$$

$\beta_j(\tau)$

Compare to crossed:  $df_A = a - 1$ ,  $df_B = b - 1$ ,  
 $df_{AB} = (a-1)(b-1)$

(8)

$$SS_A = \sum_i \frac{y_{i..}^2}{bn} - \frac{y_{...}^2}{N}$$

$$SS_{B(A)} = \sum_i \sum_j \frac{y_{ij.}^2}{n} - \sum_i \frac{y_{i..}^2}{bn}$$

Example: A: machine fixed a levels

B: operator random nested within machine  
 ↳ operators " "

n replicates

|                  | $\begin{matrix} a \\ F \\ i \end{matrix}$ | $\begin{matrix} b \\ R \\ j \end{matrix}$ | $\begin{matrix} n \\ R \\ k \end{matrix}$ | EMS   |
|------------------|---|---|---|---|
| $\tau_i$         | 0   | b   | n   | $bn \frac{\sum \tau_i^2}{a-1} + n\sigma_F^2 + \sigma^2$ |
| $\beta_{j(i)}$   | 1   | 1   | n   | $n\sigma_F^2 + \sigma^2$                                |
| $\epsilon_{ijk}$ | 1   | 1   | 1   | $\sigma^2$  |

(9)

F tests for A:  $\frac{MS_A}{MS_{B(A)}}$

$B(A) : \frac{MS_{B(A)}}{MSE}$

Example: y measures hardness of material

(10)

2 different alloy formulations

3 different heats for each alloy

2 ingots of each alloy

2 replications

A: alloy  $a=2$  fixed

B: heat  $b=3$  fixed nested within A

C: ingot  $c=2$  random nested with both A & B

$n=2$

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|                     | A<br>F<br>i | B<br>F<br>j | C<br>R<br>k | R<br>L | EMS  |
|---------------------|-------------|-------------|-------------|--------|--|
| $\tau_i$            | 0           | 3           | 2           | 2      | $12\sigma_{\tau_i}^2 + 2\sigma_{\gamma}^2 + \sigma^2$                  |
| $\beta_{j(i)}$      | 1           | 0           | 2           | 2      | $4 \frac{\sum \sum \beta_{j(i)}^2}{4} + 2\sigma_{\gamma}^2 + \sigma^2$ |
| $\gamma_{k(ij)}$    | 1           | 1           | 1           | 2      | $2\sigma_{\gamma}^2 + \sigma^2$  |
| $\epsilon_{(ijk)l}$ | 1           | 1           | 1           | 1      | $\sigma^2$   |

|                                |           |  |
|--------------------------------|-----------|--|
| A: $MS_A / MS_{C(AB)}$         | df = 1, 6 | C(AB): $\frac{MS_{C(AB)}}{MS_E}$<br>df = 6, 12 |
| B(A): $MS_{B(A)} / MS_{C(AB)}$ | df = 4, 6 |  |

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| Source | df |          |
|--------|----|----------|
| A      | 1  | a-1      |
| B(A)   | 4  | (b-1)a   |
| C(AB)  | 6  | (c-1)ab  |
| ERR    | 12 | (n-1)abc |
| Tot    | 23 | N-1      |

HW # 2

# 13.17

**13.17.** Consider a four-factor factorial experiment where factor  $A$  is at  $a$  levels, factor  $B$  is at  $b$  levels, factor  $C$  is at  $c$  levels, factor  $D$  is at  $d$  levels, and there are  $n$  replicates. Write down the sums of squares, the degrees of freedom, and the expected mean squares for the following cases. Assume the restricted model for all mixed models. You may use a computer package such as Minitab.

- (a)  $A$ ,  $B$ ,  $C$ , and  $D$  are fixed factors.
- (b)  $A$ ,  $B$ ,  $C$ , and  $D$  are random factors.
- (c)  $A$  is fixed and  $B$ ,  $C$ , and  $D$  are random.
- (d)  $A$  and  $B$  are fixed and  $C$  and  $D$  are random.
- (e)  $A$ ,  $B$ , and  $C$  are fixed and  $D$  is random.

Do exact tests exist for all effects? If not, propose test statistics for those effects that cannot be directly tested.