

Chp 7 Sec 3

Moment generating functions

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5-29

Defn: $M_X(t) = E[e^{tx}]$

$$M_X(0) = E[e^0] = E[1] = 1$$

$$\frac{d}{dt} M_X(t) = \frac{d}{dt} E[e^{tx}]$$

$$= E\left[\frac{d}{dt} e^{tx}\right]$$

$$= E[e^{tx} x]$$

$$\left. \frac{d}{dt} M_X(t) \right|_{t=0} = E[X]$$

$$\frac{d^2}{dt^2} M_X(t) = E\left[\frac{d^2}{dt^2} e^{tx}\right]$$

$$= E[e^{tx} x^2]$$

$$\left. \frac{d^2}{dt^2} M_X(t) \right|_{t=0} = E[X^2]$$

②

Find the moment generating function for the exponential distribution.

$$f(x) = \frac{1}{\beta} e^{-x/\beta} \quad x > 0$$

$$M_X(t) = E[e^{tx}] = \int_0^{\infty} e^{tx} \frac{1}{\beta} e^{-x/\beta} dx$$

$$= \frac{1}{\beta} \int_0^{\infty} e^{-x(\frac{1}{\beta}-t)} dx \quad (3)$$

$$= \frac{1}{\beta} \frac{e^{-x(\frac{1}{\beta}-t)}}{-(\frac{1}{\beta}-t)} \Big|_0^{\infty}$$

$$= 0 - \frac{1}{\beta} \frac{1}{-(\frac{1}{\beta}-t)} = \frac{1}{1-\beta t}$$

$$[\text{provided } \frac{1}{\beta}-t > 0 \text{ or } t < \frac{1}{\beta}]$$

$$\text{Check: } M_x(0) = 1 \quad \checkmark$$

$$M_x'(t) = -(1-\beta t)^{-2}(-\beta) = \beta(1-\beta t)^{-2}$$

$$M_x'(0) = \beta$$

$$\begin{aligned} M_x''(t) &= \beta(-2)(1-\beta t)^{-3}(-\beta) \\ &= 2\beta^2(1-\beta t)^{-3} \end{aligned} \quad (4)$$

$$M_x''(0) = 2\beta^2 = E[X^2]$$

$$\sigma^2 = 2\beta^2 - \beta^2 = \beta^2$$

Find $M_x(t)$ for the binomial distribution

$$\begin{aligned} M_x(t) &= E[e^{tx}] \\ &= \sum_{x=0}^n e^{tx} \binom{n}{x} p^x q^{n-x} \\ &= \sum_{x=0}^n \binom{n}{x} (pe^t)^x q^{n-x} \end{aligned}$$

$$= (pe^t + q)^n \sum_{x=0}^n \binom{n}{x} \underbrace{\left(\frac{pe^t}{pe^t + q}\right)^x}_{p^x} \underbrace{\left(\frac{q}{pe^t + q}\right)^{n-x}}_{q^{n-x}} \quad (5)$$

$$= (pe^t + q)^n$$

Check: $M_x(0) = 1 \quad \checkmark$

$$\begin{aligned} M_x'(t) &= n(pe^t + q)^{n-1} pe^t \\ &= npe^t (pe^t + q)^{n-1} \end{aligned}$$

$$M_x'(0) = np$$

$$M_x''(t) = np \left[e^t (n-1)(pe^t + q)^{n-2} pe^t + (pe^t + q)^{n-1} e^t \right]$$

$$\begin{aligned} M_x''(0) &= np[(n-1)p + 1] \quad (6) \\ &= n(n-1)p^2 + np = E[X^2] \end{aligned}$$

$$\begin{aligned} \sigma^2 &= n(n-1)p^2 + np - n^2p^2 \\ &= -np^2 + np \\ &= np(-p + 1) \\ &= npq \end{aligned}$$

Find the m.g.f. for the normal distribution

$$\begin{aligned} M_x(t) &= E[e^{tx}] \\ &= \int_{-\infty}^{\infty} e^{tx} \cdot \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx \end{aligned}$$

$$\begin{aligned}
&= \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{tx - \frac{1}{2\sigma^2}(x^2 - 2\mu x + \mu^2)} dx \quad (7) \\
&= \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2}(x^2 - 2\mu x + \mu^2 - 2\sigma^2 tx)} dx \\
&= \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2}(x^2 - 2(\mu + \sigma^2 t)x + \mu^2)} dx \\
&= \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2} \left[x^2 - 2(\mu + \sigma^2 t)x + (\mu + \sigma^2 t)^2 - (\mu + \sigma^2 t)^2 + \mu^2 \right]} dx \\
&= e^{-\frac{1}{2\sigma^2} [-(\mu + \sigma^2 t)^2 + \mu^2]} \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2} (x - (\mu + \sigma^2 t))^2} dx}_1
\end{aligned}$$

$$\begin{aligned}
&= e^{-\frac{1}{2\sigma^2} [-\mu^2 - 2\mu\sigma^2 t - \sigma^4 t^2 + \mu^2]} \quad (8) \\
&= e^{-\frac{1}{2\sigma^2} [-\sigma^2 (2\mu t + \sigma^2 t^2)]} \\
&= e^{\frac{1}{2} (2\mu t + \sigma^2 t^2)} \\
&= e^{\mu t + \frac{1}{2} \sigma^2 t^2}
\end{aligned}$$

Notice the m.g.f. for the standard normal distribution is

$$M_X(t) = e^{\frac{1}{2} t^2}$$

Properties of the m.g.f.

Start with a random variable X .

$$\text{Let } Y = X + c$$

$$\begin{aligned} M_Y(t) &= E[e^{tY}] \\ &= E[e^{t(X+c)}] \\ &= E[e^{tX} \cdot e^{tc}] \\ &= e^{tc} E[e^{tX}] \\ &= e^{tc} M_X(t) \end{aligned}$$

⑨

$$\text{Let } Y = aX$$

$$\begin{aligned} M_Y(t) &= E[e^{tY}] \\ &= E[e^{taX}] \\ &= M_X(ta) \end{aligned}$$

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Suppose X_1 and X_2 are independent random variables.

$$\text{Let } Y = X_1 + X_2$$

$$\begin{aligned} M_Y(t) &= E[e^{tY}] \\ &= E[e^{t(X_1 + X_2)}] \end{aligned}$$

$$= E[e^{tx_1} e^{tx_2}]$$

(11)

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{tx_1} e^{tx_2} f(x_1, x_2) dx_1 dx_2$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{tx_1} e^{tx_2} \underbrace{f(x_1, x_2)}_{\text{independence}} dx_1 dx_2$$

$$= \int_{-\infty}^{\infty} \left[e^{tx_2} h(x_2) \underbrace{\int_{-\infty}^{\infty} e^{tx_1} g(x_1) dx_1}_{M_{x_1}(t)} \right] dx_2$$

$$= M_{x_1}(t) M_{x_2}(t)$$

Find the m.g.f. for the gamma distribution. (12)

Think of the gamma random variable as the sum of α independent exponential random variables.

$$\text{So } M_x(t) = \left(\frac{1}{1 - \mu t} \right)^\alpha$$

Find the m.g.f. of the Poisson distribution.

$$M_x(t) = E[e^{tx}] = \sum_{x=0}^{\infty} e^{tx} \mu^x \frac{e^{-\mu}}{x!}$$

$$= \sum_{x=0}^{\infty} \frac{(\mu e^t)^x e^{-\mu}}{x!} \quad (13)$$

$$= e^{-\mu} e^{\mu e^t} \underbrace{\sum_{x=0}^{\infty} \frac{(\mu e^t)^x e^{-\mu e^t}}{x!}}_1$$

$$= e^{\mu(e^t - 1)}$$

Try the geometric, neg binom,
uniform