

For the hypergeometric distribution,

Stat 451
10-13-16
①

$$\mu = n \frac{N_1}{N}, \quad \sigma^2 = n \frac{N_1}{N} \left(1 - \frac{N_1}{N}\right) \left(\frac{N-n}{N-1}\right)$$

finite population
correction

Geometric distribution

$$\left[\begin{aligned} \text{Geometric series} &= a + ar + ar^2 + \dots = T \\ T - a &= ar + ar^2 + \dots \\ &= r(a + ar + ar^2 + \dots) = rT \\ T(1-r) &= a \quad T = \frac{a}{1-r} \end{aligned} \right]$$

②

Run a sequence of Bernoulli trials

(independent trials, 2 possible outcomes on each trial,
same probability of success " " ")

X = trial on which the 1st success occurs

$$\boxed{X = 1, 2, 3, \dots}$$
$$p(x) = q^{x-1} p$$

Check that the probs

sum to 1:

$$\begin{aligned} \sum_{x=1}^{\infty} q^{x-1} p &= p + pq + pq^2 + \dots \\ &= \frac{p}{1-q} = \frac{p}{p} = 1 \end{aligned}$$

$$\text{Find } \mu = E[X] = \sum_{\text{all } x} x p(x)$$

$$= \sum_{x=1}^{\infty} x q^{x-1} p \quad \text{let } y = x-1 \quad x = y+1$$

$$= \sum_{y=0}^{\infty} (y+1) q^y p = \underbrace{\sum_{y=0}^{\infty} y q^y p}_{(1)} + \underbrace{\sum_{y=0}^{\infty} q^y p}_{(2)}$$

$$(1): \sum_{y=1}^{\infty} y q^y p = q \sum_{y=1}^{\infty} y q^{y-1} p = q \mu$$

$$(2): p + pq + pq^2 + \dots = 1$$

$$\mu = q\mu + 1$$

$$\mu - q\mu = 1$$

$$\mu = \frac{1}{1-q} = \frac{1}{p}$$

To find σ^2 , find $E[X^2] - \mu^2$ (4)

↑ know this already

$$E[X^2] = \sum_{x=1}^{\infty} x^2 q^{x-1} p \quad \text{let } y = x-1 \text{ and proceed as before}$$

Answer: $\sigma^2 = \frac{q}{p^2}$

Negative binomial (Pascal) distribution

Same setup as the geometric distribution,
but X = trial on which the r^{th} success occurs

$$X = r, r+1, r+2, \dots$$

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$$P(X) = P(r^{\text{th}} \text{ success happens on trial } \# X)$$

$$= P(r-1 \text{ successes in the } 1^{\text{st}} X-1 \text{ trials} \cap \text{success on trial } X)$$

$$= \binom{X-1}{r-1} p^{r-1} q^{(X-1)-(r-1)} \cdot p$$

$$P(X) = \binom{X-1}{r-1} p^r q^{X-r}, \quad X = r, r+1, \dots$$

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$$X = X_1 + X_2 + \dots + X_r$$

↑
trial
on which
1st success
occurred

↑
trial
on which
2nd occurs,
starting with
new count

Each X_i has a geometric distribution, & they are all independent.

$$\begin{aligned} \text{So } \mu = E[X] &= E[X_1] + \dots + E[X_r] \\ &= \frac{r}{p} \end{aligned}$$

$$\sigma^2 = \text{Var}[X] = \text{Var}[X_1] + \dots + \text{Var}[X_r] = \frac{r q}{p^2}$$

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Example: Each attack has a 20% success rate

It takes 3 successes to eliminate the threat. (Must assume independence)

✓ Find the prob. distr. Find the expected value.

$$X \sim NB(r=3, p=.2)$$

$$p(x) = \binom{x-1}{2} (.2)^3 (.8)^{x-3}, \quad x=3, 4, \dots$$

$$\mu = \frac{r}{p} = \frac{3}{.2} = 15$$

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Find the probability that the 3rd success occurs on trial 10

$$p(10) = \binom{9}{2} (.2)^3 (.8)^7 = .0604$$

Find the prob. that the 3rd success occurs on or before trial 20.

$$P(X \leq 20) = p(3) + p(4) + \dots + p(20)$$

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$$\lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n$$

$$\text{let } y = \left(1 + \frac{x}{n}\right)^n$$

$$\ln y = n \ln\left(1 + \frac{x}{n}\right) = \frac{\ln\left(1 + \frac{x}{n}\right)}{\frac{1}{n}}$$

$$\lim_{n \rightarrow \infty} \ln y = \lim_{n \rightarrow \infty} \frac{\ln\left(1 + \frac{x}{n}\right)}{\frac{1}{n}} = \lim_{n \rightarrow \infty} \frac{\frac{1}{1 + \frac{x}{n}} \cdot \left(-\frac{x}{n^2}\right)}{\left(-\frac{1}{n^2}\right)}$$

by L'Hôpital

$$= \lim_{n \rightarrow \infty} \left(\frac{x}{1 + \frac{x}{n}} \right) = x$$

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$$\lim_{n \rightarrow \infty} \ln y = x$$

"

$$\ln(\lim y) \quad \text{because } \ln y \text{ is continuous}$$

$$\therefore \lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n = e^x$$

Start with the binomial distribution

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$$p(x) = \binom{n}{x} p^x q^{n-x}$$

Take the limit as $n \rightarrow \infty$ and $p \rightarrow 0$

and hold np constant

$$\text{Set } np = \lambda \quad \text{so } p = \frac{\lambda}{n}$$

$$p(x) = \binom{n}{x} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x}$$

$$\lim_{n \rightarrow \infty} p(x) = \lim_{n \rightarrow \infty} \frac{n!}{x!(n-x)!} \frac{\lambda^x}{n^x} \frac{\left(1 - \frac{\lambda}{n}\right)^n}{\left(1 - \frac{\lambda}{n}\right)^x}$$

$$= \frac{\lambda^x}{x!} \lim_{n \rightarrow \infty} \frac{n(n-1)\dots(n-x+1)}{\underbrace{n \cdot n \dots n}_{1(1-\frac{1}{n}) \dots (1-\frac{x-1}{n})}} \frac{\left(1 - \frac{\lambda}{n}\right)^n}{\left(1 - \frac{\lambda}{n}\right)^x}$$

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$$= \frac{\lambda^x}{x!} e^{-\lambda}$$

$$p(x) = \frac{\lambda^x e^{-\lambda}}{x!} \text{ is the Poisson distribution}$$

$$X = 0, 1, 2, \dots$$

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$$\text{Check: } \sum_{k=0}^{\infty} \frac{\lambda^k e^{-\lambda}}{k!} = e^{-\lambda} \underbrace{\left[1 + \lambda + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} + \dots \right]}_{\text{Taylor series for } e^{\lambda}}$$

$$= 1$$

$\mu = E[X] = \lambda$ because $np = \lambda$ was held constant

$$\sigma^2 = \lambda \qquad npq = \lambda \left(1 - \frac{\lambda}{n}\right)$$

Poisson process or Poisson experiment

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- observe a continuous process over a fixed amount of time or space
- X counts the # of occurrences of a certain type
- Observations in non-overlapping intervals are independent
- The probability of an occurrence in a given interval is proportional to the length of the interval
- In a very small interval, the prob. of more than 1 occurrence is approx zero

(15)

Examp. A fire station gets an average of 2.83 calls per day.

On a particular day, find the prob that it gets 4 or fewer calls.

$$X \sim \text{Poisson}(\lambda = 2.83)$$

$$p(x) = \frac{2.83^x e^{-2.83}}{x!}$$

$$\begin{aligned} P(X \leq 4) &= p(0) + p(1) + \dots + p(4) \\ &= e^{-2.83} \left(1 + 2.83 + \frac{2.83^2}{2} + \frac{2.83^3}{6} + \frac{2.83^4}{24} \right) \\ &= .8429 \end{aligned}$$

(16)

HW #3 due Oct 20

p. 89 #9, 12

p. 97 #10

p. 122 #4

p. 145 #5, 7, 13, 16