

Gamma example:

We have a Poisson process with an average
of 5 occurrences per minute. ($\lambda=5$)

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Find the probability that the waiting time
for the 3rd occurrence is more than
30 seconds.

T = waiting time for 3rd occurrence.

$T \sim \text{Gamma}(\alpha = 3, \beta = \frac{1}{\lambda} = \frac{1}{5})$

$$P(T > .5) = \int_{.5}^{\infty} \frac{1}{(\frac{1}{5})^3 2!} x^2 e^{-5x} dx$$

②

$$\boxed{P(T > t) = P(X \leq \alpha - 1)}$$

\uparrow \uparrow
 $\text{Gamma}(\alpha, \beta = \frac{1}{\lambda})$ $\text{Poisson}(\lambda t)$

$$\begin{aligned} P(T > .5) &= P(X \leq 2) \quad \mu = 5(.5) = 2.5 \\ &= p(0) + p(1) + p(2) \\ &= \left(\frac{2.5^0}{0!} + \frac{2.5^1}{1!} + \frac{2.5^2}{2!} \right) e^{-2.5} \end{aligned}$$

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Same example, but find the probability that
the waiting for the 4th occurrence is
anywhere between 1 minute and 2 minutes.

$$T \sim \text{Gamma}(\alpha = 4, \beta = \frac{1}{\lambda} = \frac{1}{5})$$

$$P(1 < T < 2) = \int_1^2 \frac{1}{(\frac{1}{5})^4 3!} x^3 e^{-5x} dx$$

$$= P(T > 1) - P(T > 2)$$

$$P(T > 1) = P(X \leq 3) \quad \text{where } X \sim \text{Poisson}$$

$$\begin{aligned}\mu &= \lambda t \\ &= 5 \cdot 1 = 5\end{aligned}$$

$$= p(0) + p(1) + p(2) + p(3)$$

$$= e^{-5} \left(\frac{5^0}{0!} + \frac{5^1}{1!} + \frac{5^2}{2!} + \frac{5^3}{3!} \right)$$

$$P(T > 2) = P(X \leq 3) \quad \text{where } X \sim \text{Poisson}$$

$$= e^{-10} \left(\frac{10^0}{0!} + \frac{10^1}{1!} + \frac{10^2}{2!} + \frac{10^3}{3!} \right) \quad \begin{aligned}\mu &= \lambda \cdot t = 5 \cdot 2 \\ &= 10\end{aligned}$$

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For the gamma distribution, $\mu = \alpha\beta$
 $\sigma^2 = \alpha\beta^2$

Another special case of the gamma distribution:

Set $\alpha = \frac{v}{2}$, where v is a positive integer

and $\beta = 2$.

$$f(x) = \frac{1}{2^{\frac{v}{2}} \Gamma(\frac{v}{2})} x^{\frac{v}{2}-1} e^{-\frac{1}{2}x}$$

This is the Chi-squared density with parameter v .

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Chp7 Sec.3

Moment-generating function

Definition: $M_X(t) = E[e^{tx}]$

Properties: $M_X(0) = E[e^0] = E[1] = 1$

$$\frac{d}{dt} M_X(t) = \frac{d}{dt} E[e^{tx}]$$

$$= E\left[\frac{d}{dt} e^{tx}\right] = E[e^{tx} \cdot x]$$

$$\left. \frac{d}{dt} M_X(t) \right|_{t=0} = E[X]$$

$$\begin{aligned}\frac{d^2}{dt^2} M_X(t) &= E\left[\frac{d^2}{dt^2} e^{tx}\right] \\ &= E[e^{tx} \cdot X^2]\end{aligned}$$

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$$\left. \frac{d^2}{dt^2} M_X(t) \right|_{t=0} = E[X^2]$$

Find the mgf of the exponential distribution.

$$f(x) = \frac{1}{\beta} e^{-x/\beta} \quad x > 0$$

$$M_X(t) = E[e^{tx}] = \int_0^\infty e^{tx} f(x) dx$$

$$M_X(t) = \int_0^\infty e^{tx} \frac{1}{\beta} e^{-x/\beta} dx$$

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$$= \frac{1}{\beta} \int_0^\infty e^{-x(\frac{1}{\beta} - t)} dx$$

$$= \frac{1}{\beta} \left[\frac{e^{-x(\frac{1}{\beta} - t)}}{-(\frac{1}{\beta} - t)} \right]_0^\infty$$

$$= \frac{1}{\beta} \left[0 - \frac{1}{-(\frac{1}{\beta} - t)} \right] = \frac{1}{1 - \beta t}$$

$$\text{provided } \frac{1}{\beta} - t > 0 \\ t < \frac{1}{\beta}$$

$$\text{Check } M_x(t) = \frac{1}{1-\beta t} = 1 \quad \checkmark$$

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$$\begin{aligned}\frac{d}{dt} M_x(t) &= \frac{d}{dt} (1-\beta t)^{-1} = - (1-\beta t)^{-2} (-\beta) \\ &= \beta (1-\beta t)^{-2}\end{aligned}$$

$$\left. \frac{d}{dt} M_x(t) \right|_{t=0} = \beta = E[X]$$

$$\frac{d^2}{dt^2} M_x(t) = \beta (-2)(1-\beta t)^{-3} (-\beta) = 2\beta^2 (1-\beta t)^{-3}$$

$$\left. \frac{d^2}{dt^2} M_x(t) \right|_{t=0} = 2\beta^2 = E[X^2]$$

$$\sigma^2 = E[X^2] - (E[X])^2 = 2\beta^2 - \beta^2 = \beta^2 \quad (10)$$

mgt for binomial:

$$M_x(t) = E[e^{tx}] = \sum_{\text{all } x} e^{tx} p(x)$$

$$= \sum_{x=0}^n e^{tx} \binom{n}{x} p^x q^{n-x}$$

$$= \sum_{x=0}^n \binom{n}{x} (\rho e^t)^x q^{n-x}$$

$$= (pe^t + q)^n \sum_{x=0}^n \binom{n}{x} \left(\frac{pe^t}{pe^t + q}\right)^x \left(\frac{q}{pe^t + q}\right)^{n-x} \quad (11)$$

$$= (pe^t + q)^n \underbrace{\sum_{x=0}^n \binom{n}{x} (p')^x (q')^{n-x}}_1$$

$$M_x(t) = (pe^t + q)^n$$

$$M_x(0) = (p+q)^n = 1 \quad \checkmark$$

$$M'_x(t) = n(pe^t + q)^{n-1} \cdot pe^t$$

$$M'_x(0) = np = E[X]$$

$$M''_x(t) = np \left[(pe^t + q)^{n-2} e^t + e^t(n-1)(pe^t + q)^{n-2} pe^t \right]$$

$$M''_x(0) = np [1 + (n-1)p] \quad (12)$$

$$= np + n(n-1)p^2 = E[X^2]$$

$$\sigma^2 = np + n(n-1)p^2 - (np)^2$$

$$= np + n^2 p^2 - np^2 - n^2 p^2$$

$$= np(1-p) = npq$$

Mgf for gamma:

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$$\begin{aligned}
 M_X(t) &= E[e^{tx}] = \int_0^\infty e^{tx} \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-x/\beta} dx \\
 &= \int_0^\infty \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-x(\frac{1}{\beta}-t)} dx \\
 &= \frac{(\frac{1}{\beta}-t)^{-\alpha}}{\beta^\alpha} \underbrace{\int_0^\infty \frac{1}{[(\frac{1}{\beta}-t)^{-1}]^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-\frac{x}{(\frac{1}{\beta}-t)}} dx}_1
 \end{aligned}$$

$$\begin{aligned}
 M_X(t) &= \frac{(\frac{1}{\beta}-t)^{-\alpha}}{\beta^\alpha} = \frac{1}{[\beta(\frac{1}{\beta}-t)]^\alpha} \\
 &= (1-\beta t)^{-\alpha}
 \end{aligned}$$

Check: $M_X(0) = 1 \quad \checkmark$

$$\begin{aligned}
 M'_X(t) &= -\alpha (1-\beta t)^{-\alpha-1} (-\beta) \\
 &= \alpha \beta (1-\beta t)^{-\alpha-1}
 \end{aligned}$$

$$M'_X(0) = \alpha \beta = E[X]$$

$$M_X''(t) = \alpha\beta(-\alpha-1)(1-\beta t)^{-\alpha-2}(-\beta) \quad (15)$$

$$= \alpha\beta^2(\alpha+1)(1-\beta t)^{-\alpha-2}$$

$$M_X''(0) = -\alpha\beta^2(\alpha+1) = \alpha^2\beta^2 + \alpha\beta^2$$

$$= E[X^2]$$

$$\sigma^2 = \alpha^2\beta^2 + \alpha\beta^2 - (\alpha\beta)^2 = \alpha\beta^2$$