Statistical Significance of the Seismic Coupling Coefficient

by Robert McCaffrey

Abstract Variations in the seismic coupling coefficient (Z), the ratio of the observed to expected seismic moment release rates, at subduction zones are commonly used to make inferences about the physics of subduction. Taking into account the power-law form of the earthquake frequency–moment relation, it is shown that the observed distribution of Z at the world’s subduction zones based on the twentieth century’s earthquakes can be matched by a single value of $Z = 0.3$ if, as is likely the case, the 90-yr observation time is less than or comparable to the repeat time of the largest possible earthquake. It is shown that because our reliable earthquake record is short, global variations in Z, based on seismicity alone, are poorly resolved and cannot be used to distinguish among subduction models.

Introduction

The theory of plate tectonics has, among other things, allowed us to make predictions about the long-term rate of seismic moment release on faults. Early studies utilizing this predictive capacity found that over time some faults generate roughly the correct amount of seismic moment in earthquakes, but others do not (Brunner, 1968; Davies and Bruner, 1971). Peterson and Seno (1964) calculated the ratios of the observed to the expected moment rates for most of the world’s subduction zones and interpreted these ratios as measures of the coupling on the thrust fault. They went on to relate the variations in this moment ratio to other subduction parameters in order to gain insight into the physics of subduction. Scholz (1990) renamed this ratio the seismic coupling coefficient $Z$, and Pacheco et al. (1993) recalculated values of $Z$ for the world’s subduction zones using the recently revised catalog of Pacheco and Sykes (1992), which includes earthquakes of $M_o \geq 7$ in this century. Scholz and Campos (1995) amended $Z$ for some of the world’s subduction zones by including large earthquakes from the nineteenth century and interpreted variations in it as support for their slab anchor model of subduction.

Brunner (1968), whose purpose was to estimate seismic slip rates and not $Z$, recognized that accurate estimates required a long seismic history due to the random nature of earthquake occurrence. Most other workers who worked with such data also recognized the inherent dangers in making generalizations regarding the long-term seismic behavior of a fault based on the short record. Nevertheless, global variations in $Z$ continue to be used to make inferences about the physics of subduction. The purpose of this article is to test the randomness of our observed $Z$ values that are based on, at best 90 yr and at worst 30 yr, of reliable earthquake moments. It is found that the range of observed $Z$ values could be generated by subduction zones all with the same actual $Z$ and seismicity distributions. Only the smallest values ($Z < 0.1$), which are difficult to get at random from a large value of $Z$, may indicate fundamentally different behavior. However, Z is possible to get the full range of observed $Z$ from a small actual $Z$ if the observation time is less than about the repeat time of the largest earthquake. Because such repeat times are likely hundreds of years, today’s $Z$ variations probably do not provide useful constraints on the relative merits of subduction models.

Method and Data

We start with the distribution of earthquake moments of the form $N(M_o) = a M_o^b$, where $N$ is the number of earthquakes with a seismic moment greater than or equal to $M_o$, and $a$ and $b$ are constants that depend on the frequency–magnitude and moment–magnitude relationships (Wyss, 1973; Richter, 1958). $b$ is related to the more common $b$-value, which is the slope of the log (frequency–magnitude) relationship, by $b = 2b/3$. $M_o$ and $b$ will be used interchangeably, noting the relation $M_o = 2/3 \log M_o - 6.0$, where $M_o$ is in Newton-meters (Hanks and Kanamori, 1979).

Based on earthquake self-similarity, $b$ should equal 2/3 for earthquakes whose slip dimensions are unbounded by the edges of the fault surface and 1 for the larger events whose slip is limited in one dimension (Rundle, 1989). From the global earthquake catalog of 1900 to 1990, Pacheco et al. (1992) estimated that $b = 0.6$ for $7.0 \leq M_o \leq 7.5$ and $b = 0.87$ for $7.5 \leq M_o \leq 8.8$, generally in agreement with Rundle’s prediction (Pacheco et al. actually estimated $b$). They found that earthquakes of $M_o > 8.8$ are more frequent than predicted by their $b$-values, which they attributed to
insufficient sampling of the past 90 yr for the truly great earthquakes (their data are plotted in Fig. 1a). However, the
disagreement of the $M_e > 8.8$ earthquakes that have dip
areas of 100 km or more, with the predictions of self-
similarity, may alternatively be due to breakdown of one of the
self-similarity assumptions, that is, that the earthquakes do
not interact in space or time (Rundle, 1989). In this case, $\beta$
for $M_e > 8.8$ may not be the same as that observed for the
range 7.5 $\leq M_e \leq 8.8$.

$\beta$ controls the ratio of the moment of the largest event
to the total moment (Fig. 1b). As $\beta$ increases, more of the
total moment is contained in the smaller, more frequent
events. Hence, larger $\beta$ should result in a more reliable es-
imate of $\beta$ at any given time. At most of the world’s sub-
duction zones, the ratio of the largest earthquake moment
(in the twelfth century) to total moment is about 20% to
40% (Petersen and Sero, 1984), suggesting that $\beta$ is between
0.6 and 0.8 (Fig. 1). In the tests, we use the $M_{eq}$ distrib-
utions shown in Figure 1a, some of which are observed,
some theoretical, some use a single $\beta$, and others use separate
$\beta$ for small and large earthquakes.

To determine $\alpha$, we say that in time $T$ there will be only
one earthquake of moment $M_{eq}$ or larger, so that $N(M_{eq})$
$= 1$, or $\alpha = (M_{eq})^{M_{eq}}$. $T$ is the "cycle time" defined by Run-
dle (1989) as the time span during which each earthquake
in the distribution has a probability of 1/T. For the sake of
the computation, we specify a lower bound on the earth-
quake magnitudes considered ($M_{eq}$) such that all earthquakes with mo-
ments less than it contribute less than 1% to the total seismic moment. Finally, we say that the one earthquake of moment
$M_{eq}$ or larger has a moment of exactly $M_{eq}$ (its minimum
possible value) because any larger moment will increase the
ratio of the largest to total moment, leading to poorer resol-
ution of $\beta$. In other words, by choosing the minimum mo-
ment for this largest earthquake within the bounds $M_{eq}$ $\leq M_e$ $\leq$
$M_{eq} + \Delta M_e$, we maximize the resolution of $\beta$. The number of
earthquakes within a range of moments, say from $M_0 = \Delta M_0$
to $M_0 + \Delta M_0$, is then

$$n(M_0) = \{N(M_0 - \Delta M_0) - N(M_0 + \Delta M_0)\}$$
when $M_{eq}$ $\leq M_0 \leq M_{eq} + \Delta M_0$, \hspace{1cm} (1)

$$n(M_0) = 1 \hspace{1cm} \text{when} \hspace{1cm} M_0 = M_{eq}$$

$$n(M_0) = 0 \hspace{1cm} \text{when} \hspace{1cm} M_0 < M_{eq} \hspace{1cm} \text{or} \hspace{1cm} M_0 > M_{eq}$$

(Steps $\Delta M_0$ are actually taken as steps of 0.15 in the
exponent of the moment to get steps of 0.1 in the magnitude.)

From the definition of $T$, the size of the sizes of earthquakes
has a probability of occurring in proportion to the number
expected; that is, the probability of an earthquake $M_e$ oc-
curring within a time interval $\Delta T$ is

$$P(M_e) = n(M_0)/T$$

where $T$ is in years and $k$ is the number of time intervals per
year. Accordingly, $\beta$ and $M_{eq}$ define a probability density

Figure 1. (a) Numbers of earthquakes ($N(M_0)$)
greater than $M_0$ as a function of $M_0$ for the test cases
presented here. Numbers on lines refer to distributions
used in test cases shown in Figure 2 as follows: (1)
$M_{eq} = 9.5$, $b = 1$ (Figs. 2a and 2b); (2) $M_{eq} = 8.6$, $b = 0.9$ for
$M_e \leq 7.5$, $b = 1.3$ for $7.5 \leq M_e \leq 8.6$ (Fig. 2c); (3) $M_{eq} = 9.5$, $b = 0.75$ (Fig. 2d); (4)
$M_{eq} = 9.5$, $b = 1.0$ for $M_e \leq 7.5$, $b = 1.5$ for $7.5 \leq M_e \leq 9.5$
$M_e \leq 9.5$ (Fig. 2h); (5) $M_{eq} = 9.5$, $b = 0.9$ for
$M_e \leq 7.5$, $b = 1.3$ for $7.5 \leq M_e \leq 9.5$ (Figs. 2f through
2j); and (6) dotted $M_{eq} = 9.5$, $N(M_0)$ is ob-
served distribution (1900 to 1990) in the range 7.0 $\leq$
$M_e \leq 9.5$ and $b = 1$ for $M_e < 7.0$, from Pacheco et
al. (1992). Moment is plotted here in terms of $M_e$
instead of $M_0$, so the slopes are $-b$ instead of $-\beta \beta$
$= 26.3$). (b) Solid line shows the ratio (in percent) of the
seismic moment of the largest earthquake to the total
moment, as a function of $\beta$. Histograms show the
distribution of these percentages in subduction zones,
taken from Petersen and Sero (1984). The me-
dian is around 35%, corresponding to a $\beta$ of around
0.6 to 0.8.
function as a function of moment for the earthquakes during time $T$. For example, if $b = 2.3$ and $M_{W} = 9.5$, the total number of expected events of $M_{W} \geq 5.0$ is $31,623$. If $T = 200$ yr and $k = 365$ (days), then the probability of an earthquake of $9.5 \geq M_{W} \geq 5.0$ occurring on any particular day is $31.623 \times 365 \times 200$. By definition, one earthquake of $M_{W}^{*}$ occurs in time $T$, so its probability is $1/T$. The total expected moment in time $T$ is $M_{W} = \Sigma M_{W} \epsilon(M_{W})$, (summation from $M_{W} = 5.0$ to $M_{W}^{*}$), and the expected moment rate is then $M_{W}/T$.

The procedure used to calculate the seismic coupling coefficient at time $t$ (in years) is to sample the probability density function (2) once during each time step from time 0 to time 4t. Because the probability of an earthquake of $M_{W}$ occurring in the time period 0 to 4t is $b \epsilon(M_{W})4t$, the time step used does not matter in practice as long as the number of time steps in $T$ is greater than the number of earthquakes in $T$. Each time step will either have an earthquake of $M_{W} \geq 5.0$ or no earthquake depending on a random sample of the probability density function. The summed seismic moment for all events that occur up to time $t$ is $M_{W}$. The apparent seismic coupling coefficient at time $t$ is then

$$X(t) = \beta M_{W}^{*}(t)/M_{W}(t),$$

where $X_{0}$ is the test value of $X$, which can range from 0 to 1, and the denominator is the expected moment in the time 0 to $t$. The random sampling procedure is run 1000 times each for a duration of $2T$ ($T = 365$) for various combinations of $b, X_{0}, T$, and $M_{W}^{*}$. For each trial, $X$ is calculated at $t = T/6, T/3, T/2, T, T+T$, and $2T$ and plotted as frequency distributions (Fig. 2).

Before presenting the results, a discussion of the meaning of $T$ is warranted because the distributions of $X$ depend on $T$. Also, because our observed values of $X$ are based on seismicity of the twentieth century only, we must estimate what fraction of $T$ this 90-yr sample represents. The term “recurrence time” commonly refers to the time between the large earthquakes on a fault, and these are often those larger than about magnitude 8. If the largest expected earthquake in time $T$ is a 9.5, then there should also be several earthquake of $M_{W} \geq 8$ (Fig. 1a). Hence, $T$ is not the recurrence time, as popularly used in the sense of being the interval between $M_{W} \geq 8$ earthquakes. The Nankai Trough has had about 12 earthquakes of $M_{W} \geq 8$ since 684 A.D. (Scholz, 1990), which would suggest that $T$ is more than 2000 yr if the largest earthquake there is a 9.5. However, the largest known earthquakes at the Nankai Trough are about $M_{W} = 8.6$ (in 1498 and 887 A.D.). Using $M_{W}^{*} = 8.6$ and $\beta = 0.87$ predicts six earthquakes of $M_{W} \geq 8$ during the time $T$ (Fig. 1a). Taking the magnitudes at face value, two earthquakes of the maximum size 8.6 and 12 of $M_{W} \geq 8$ indicates that two complete cycles have elapsed since 684 A.D. Hence, $T$ for the Nankai Trough is probably close to 650 yr, and the twentieth century is represented by $T/7$. For most subduction zones, we probably do not have a long enough record to evaluate $T$, and trying to assign $M_{W}^{*}$ for a trench based on the twentieth century’s seismicity involves circular reasoning because it assumes that the biggest one has occurred. Nevertheless, we will see that $T$ would have to be as short as a few decades for the 90-yr record to provide reliable estimates of $X$.

Using earthquake statistics for the 90-yr record, we can estimate a minimum global value of $T$. A global value of $M_{W}^{*} = 9.5$ is used because the 1960 Chile earthquake is the largest reliably known to us in recorded history. How- ever, observed frequency-magnitude relationship for $7.5 \leq M_{W} \leq 8.8$ earthquakes in the twentieth century predict a single event of $M_{W} = 9.3$ and not 9.5 (Pacheco et al., 1992) (dashed line in Fig. 1a). Their straight line of log $\nu(M_{W})$ versus $M_{W}$ can be shifted to intersect $N = 1$ at $M_{W} = 9.5$ by multiplying $N$ by 2, suggesting that, globally, for $M_{W}^{*} = 9.5$, $T = 180$ yr, or at present, $T = 72$. For most of the test cases, I use $M_{W}^{*} = 2 \times 10^{5}$ Nm ($M_{W}^{*} = 9.5$), but it is shown, and self-similarity predicts, that $X$ does not depend on $M_{W}^{*}$ used.

The calculated distributions of $X$ will be compared to the distribution of observed values for 23 of the world’s subduction zones (Pacheco et al., 1993). These $X$’s are based on 90 yr of data from the catalog of Pacheco and Sykes (1992), and I do not take into consideration the uncertainties in them. Nor do I consider the amendments of Scholz and Cumpeos (1995), who included earthquakes from the nineteenth century at some subduction zones, because the nineteenth century earthquakes cause heterogeneity in the data set, because the uncertainty in moments for them must be at least an order of magnitude, and because the revised $X$ are too few to impact my conclusions.

Results and Discussion

Figure 2 shows calculated distributions of $X$ at six times for each of several combinations of $T, b, X_{0}$, and $M_{W}^{*}$. The top distribution in each panel is the average of the lower six and can be thought to represent the distribution of $X$ at a fixed time from several subduction zones that have different values of $T$. The calculated distributions overall show poor resolution of $X$ at small $t$. Not only are the variances large, but in most cases, the most probable values (the peaks) are quite different from the true values (shown by dashed lines). Only when $t$ gets to be a large fraction of, or larger than, $T$ does the calculated distribution have its peak at $X_{0}$. Even at $t = 2T$, the distributions are quite broad compared to the useful range of $X$, which is between 0 and 1.

What is not evident from Figure 2 is that at any time the mean of $X$ is in fact equal to $X_{0}$. This is expected because the outcome at each time step is independent of all other time steps. The number of independent samples at any time $t$ is the number of trials $N_{t}$ times the number of time steps $k$, where $k$ is the number of time steps per year. In terms of the average calculated $X$, this is the same as using $k$ trials.
Figure 2. Calculated distributions of $x$ at different normalized times and for varying values of $b$, $T$, and $M_{850}$. Each distribution is labeled by its time as a fraction of $T$. Top: distribution in each panel (labeled "All") in the average of the six lower ones and simulates the effect of having subduction zones with variable $T$ values. True $x_0$ are shown by vertical dashed lines (in top row of panels, $x_0 = 0.5$; in middle and bottom rows, $x_0$ varies from 0.1 at left to 0.9 at right in steps of 0.2). Distributions are made in increments of 0.05 in $x$, and vertical scale is 25% as marked. Darker curve at bottom is the observed global distribution of $x$ for the twentieth century, plotted at the same scale. In Figures 2f through 2j, $b$-values estimated by Pacheco et al. (1992) we used ($b = 0.8$ for $M_{850} = 7.5$ and $b = 1.3$ for $7.5 < M_{850} < 9.5$). $M_{850} = 9.5$ is the projection of these $b$-values to the axis at $N(M_{850}) = 1$ (Fig. 1a).
and $N_t$ time-steps. However, the normalized distributions of $X$ (Fig. 2) depend on $k$ but not on $N_t$ (unless $N_t$ is too small).

For the Earth, the number of trials is $23$ (the number of subduction zones), and the average $X$ for them is $0.3$. In tests using $N_t = 23$ and $r = 90$, the average calculated $X$ is within 0.05 of the real $X$, suggesting that the global average of $X \approx 0.5$ may be robust.

The calculated distributions are compared to the 90-yr observed distribution of $X$ (at bottom), in thicker lines, in Fig. 2 taken from Pacheco et al. (1993). The similarity of Figs. 2a and 2b shows that the results depend much more on the elapsed time normalized by $T$ than on $T$ itself. The choice of $M_0^{\infty}$ has no influence on the $X$ distributions (compare Figs. 2c to 2b) because it merely shifts the earthquake probability density function along the magnitude axis and does not impact the relative probabilities of any pair of magnitudes below $M_0^{\infty}$.

The relative probabilities of the earthquake magnitudes, determined by $a$, is the major influence on $X$. Increasing $b$ has the effect of increasing the numbers of small earthquakes relative to the big ones, which results in more rapid development of a peak around the correct value of $X$ (compare Fig. 2a to 2d). Small $a$ produces fewer small events relative to the large ones, resulting in a very peaked distribution (but peaks are not at the correct value; Fig. 2d) because the large, infrequent earthquakes dominate.

The lower values of $X_0$ appear to be better resolved than the high values ($X$ varies across the bottom two rows of Fig. 2) because both the mean and standard deviation of the calculated $X$ depend directly on $X_0$. For low values of $t$, the peaks in the estimated values of $X$ do not correspond with $X_0$. Note that for $X_0 = 0.9$, the true value is not evident until $t \approx 7$ (Figs. 2c and 2d).

Figure 2 (middle row) shows tests using $N_0$, distributions taken from the twentieth-century’s earthquakes. With the constraint $X_0 = 0.3$ (Fig. 2g), a distribution similar to the observed $v_x$ is produced when $t \leq 772$. In Figures 2k through 2o, the observed magnitude distribution (dots in Fig. 1a) is used for $N_0$, in the range 7.0 $M_0 \leq 9.5$ and $b = 0.9$ for $M_0 \leq 7.0$. Again, for the low values of $X_0$ and $t$, the observed distribution of $X$ is matched quite well (Figs. 2l and 2m).

The point of this report, and shown by Figure 2, is that any particular value of $X_0$ can give rise to a large range of apparent values for $X$, especially at times that are short compared to $T$. Moreover, the observed global distribution of $X$ at subduction zones is muted by some of these randomly generated distributions. The observed distribution is best matched at $t \approx 76$ or $t \approx 772$. As suggested above, $T = 180$ yr for the whole Earth, but for individual subduction zones, it is probably hundreds of years. If so, then estimates of the variations in $X$ at subduction zones are not usefully constrained.

The inability in theory to estimate $X$ will only be compounded by the consideration of uncertainties in the $X$ estimates themselves, due largely to uncertainties in seismic moments for great and old earthquakes. About all one can say with confidence is that subduction zones without any known history of significant earthquakes probably have a low seismic coupling coefficient. However, we cannot say that subduction zones with known significant earthquakes have a large coupling coefficient. In summary, at present, we cannot rule out that most or all subduction zones have a moderate seismic coupling coefficient of around 0.3.

References


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