

# The multivariate normal distribution

Stat 571

10-17-13

①

Univariate case,  $f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$

$-\infty < x < \infty$

Multivariate case:

$$f(\vec{x}) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} e^{-\frac{1}{2}(\vec{x}-\vec{\mu})' \Sigma^{-1} (\vec{x}-\vec{\mu})}$$

$-\infty < x_i < \infty$   
for  $i=1, \dots, p$

Special case:

②

Bivariate normal ( $p=2$ )

$$\vec{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \quad \vec{\mu} = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \quad \Sigma = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{12} & \sigma_{22} \end{bmatrix}$$

$$\Sigma^{-1} = \frac{1}{\sigma_{11}\sigma_{22} - \sigma_{12}^2} \begin{bmatrix} \sigma_{22} & -\sigma_{12} \\ -\sigma_{12} & \sigma_{11} \end{bmatrix}$$

$$(\vec{x}-\vec{\mu})' \Sigma^{-1} (\vec{x}-\vec{\mu})$$

$$= [(x_1-\mu_1), (x_2-\mu_2)] \frac{1}{\sigma_{11}\sigma_{22} - \sigma_{12}^2} \begin{bmatrix} \sigma_{22} & -\sigma_{12} \\ -\sigma_{12} & \sigma_{11} \end{bmatrix} \begin{bmatrix} x_1-\mu_1 \\ x_2-\mu_2 \end{bmatrix}$$

$$= \frac{\sigma_{22}(x_1 - \mu_1)^2 - 2\sigma_{12}(x_1 - \mu_1)(x_2 - \mu_2) + \sigma_{11}(x_2 - \mu_2)^2}{\sigma_{11}\sigma_{22} - \sigma_{12}^2} \quad (3)$$

$$= \frac{\frac{(x_1 - \mu_1)^2}{\sigma_{11}} - 2\sigma_{12} \frac{(x_1 - \mu_1)}{\sigma_{11}} \frac{(x_2 - \mu_2)}{\sigma_{22}} + \frac{(x_2 - \mu_2)^2}{\sigma_{22}}}{1 - \frac{\sigma_{12}^2}{\sigma_{11}\sigma_{22}}}$$

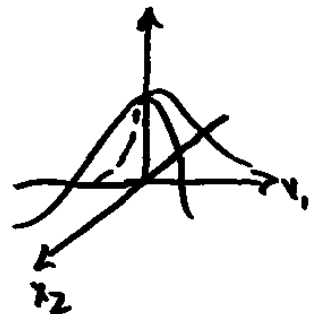
$$= \frac{1}{1 - \rho^2} \left[ \left( \frac{x_1 - \mu_1}{\sigma_1} \right)^2 - 2\rho \left( \frac{x_1 - \mu_1}{\sigma_1} \right) \left( \frac{x_2 - \mu_2}{\sigma_2} \right) + \left( \frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right]$$

$$= Q$$

$$f(x_1, x_2) = \frac{1}{2\pi |\Sigma|^{1/2}} e^{-\frac{1}{2}Q} \quad (4)$$

$$\begin{aligned} |\Sigma|^{1/2} &= \sqrt{\sigma_{11}\sigma_{22} - \sigma_{12}^2} \\ &= \sqrt{(\sigma_{11}\sigma_{22}) \left(1 - \frac{\sigma_{12}^2}{\sigma_{11}\sigma_{22}}\right)} \\ &= \sigma_1\sigma_2\sqrt{1 - \rho^2} \end{aligned}$$

$$f(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1 - \rho^2}} e^{-\frac{1}{2}Q}$$



The contours of the graph are found  
by setting  $f(x, x_0) = k$ ,

which is equivalent to setting  $Q = \text{constant}$

$$(\vec{x} - \vec{\mu})' \Sigma^{-1} (\vec{x} - \vec{\mu}) = c^2$$

These are ellipses centered at  $\vec{\mu}$ ,

and whose axes are the eigenvectors  
of  $\Sigma$ , with 1/2-lengths  $c\sqrt{\lambda_1}$  and  $c\sqrt{\lambda_2}$ ,

where  $\lambda_1$  &  $\lambda_2$  are the eigenvalues of  $\Sigma$ .

Special case of the bivariate normal:  $\rho = 0$

$$\begin{aligned} f(x_1, x_2) &= \frac{1}{2\pi \sigma_1 \sigma_2} e^{-\frac{1}{2} \left[ \left( \frac{x_1 - \mu_1}{\sigma_1} \right)^2 + \left( \frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right]} \\ &= \frac{1}{\sqrt{2\pi} \sigma_1} e^{-\frac{1}{2} \left( \frac{x_1 - \mu_1}{\sigma_1} \right)^2} \frac{1}{\sqrt{2\pi} \sigma_2} e^{-\frac{1}{2} \left( \frac{x_2 - \mu_2}{\sigma_2} \right)^2} \end{aligned}$$

So  $x_1$  and  $x_2$  are independent.

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Back to the  $\rho \neq 0$  case:

Find the marginal density of  $X_1$ .

$$f(x_1) = \int_{-\infty}^{\infty} f(x_1, x_2) dx_2$$

(7)

$$= \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} e^{-\frac{1}{2} \frac{1}{1-\rho^2} \left[ \left(\frac{x_1-\mu_1}{\sigma_1}\right)^2 - 2\rho \left(\frac{x_1-\mu_1}{\sigma_1}\right) \left(\frac{x_2-\mu_2}{\sigma_2}\right) + \left(\frac{x_2-\mu_2}{\sigma_2}\right)^2 \right]} dx_2$$

Let  $z_1 = \frac{x_1-\mu_1}{\sigma_1}$ ,  $z_2 = \frac{x_2-\mu_2}{\sigma_2}$

$\frac{dz_2}{dx_2} = \frac{1}{\sigma_2}$   $dx_2 = \sigma_2 dz_2$

$$= \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{1}{1-\rho^2} [z_1^2 - 2\rho z_1 z_2 + z_2^2]} \sigma_2 dz_2$$

(8)

$$\begin{aligned} & z_2^2 - 2\rho z_1 z_2 + z_1^2 \\ &= z_2^2 - 2\rho z_1 z_2 + \rho^2 z_1^2 + z_1^2 - \rho^2 z_1^2 \\ &= (z_2 - \rho z_1)^2 + z_1^2(1-\rho^2) \end{aligned}$$

$$\begin{aligned} \text{Now } f(x_1) &= \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{1}{1-\rho^2} [(z_2 - \rho z_1)^2 + z_1^2(1-\rho^2)]} \sigma_2 dz_2 \\ &= \frac{e^{-\frac{1}{2} z_1^2}}{2\pi\sigma_1\sqrt{1-\rho^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{1}{1-\rho^2} (z_2 - \rho z_1)^2} dz_2 \end{aligned}$$

$$= \frac{e^{-\frac{1}{2}z_1^2}}{2\pi\sigma_1\sqrt{1-p^2}} \underbrace{\int_{-\infty}^{\infty} \frac{e^{-\frac{1}{2}\left(\frac{z_2 - \rho z_1}{\sqrt{1-p^2}}\right)^2}}{\sqrt{2\pi}\sqrt{1-p^2}} dz_2}_{1} \quad (9)$$

$$= \frac{e^{-\frac{1}{2}z_1^2}}{\sqrt{2\pi}\sigma_1}$$

General result: If  $\vec{X} \sim N_p(\vec{\mu}, \Sigma)$ ,  
then  $X_i \sim N(\mu_i, \sigma_{ii})$

Theorem: If  $\vec{X} \sim N_p(\vec{\mu}, \Sigma)$ , (10)

then  $\vec{a}'\vec{X} \sim N(\vec{a}'\vec{\mu}, \vec{a}'\Sigma\vec{a})$

Theorem: If  $\vec{a}'\vec{X} \sim N(\vec{a}'\vec{\mu}, \vec{a}'\Sigma\vec{a}) \quad \underline{\underline{\forall \vec{a}}}$ ,  
then  $\vec{X} \sim N_p(\vec{\mu}, \Sigma)$ .

Theorem: If  $\vec{X}$  is  $N_p(\vec{\mu}, \Sigma)$  and  
 $A$  is a  $q \times p$  matrix of constants, then  
 $A\vec{X} \sim N_q(A\vec{\mu}, A\Sigma A')$

⑪

Hw #3 due Oct 24

4. 1,3,4,6