

Theorem: If X and Y are independent, Stat 567
2-9-17
then $E[g(X)h(Y)] = E[g(X)]E[h(Y)]$ ①

Proof: (Continuous case)

$$E[g(X)h(Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y) f(x,y) dx dy$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y) \overbrace{f_x(x)f_y(y)}^{\text{by independence}} dx dy$$

$$= \int_{-\infty}^{\infty} [h(y)f_y(y) \int_{-\infty}^{\infty} g(x)f_x(x) dx] dy$$

$$= E[g(X)] \int_{-\infty}^{\infty} h(y)f_y(y) dy$$

$$= E[g(X)] E[h(Y)].$$

②

Defn: The covariance of X & Y is

$$\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$$

Note: $E[XY - \mu_X Y - X \mu_Y + \mu_X \mu_Y]$

$$= E[XY] - \mu_X E[Y] - \mu_Y E[X] + \mu_X \mu_Y$$

$$= E[XY] - \mu_x \mu_y = E[XY] - E[X]E[Y] \quad (3)$$

Note: If X and Y are independent,

$$\text{then } E[XY] = E[X]E[Y]$$

$$\text{and } \text{Cov}(X, Y) = 0.$$

The converse is not true

Properties of the Covariance:

(4)

$$\begin{aligned} (1) \text{Cov}(X, X) &= E[X^2] - (E[X])^2 \\ &= V[X] \end{aligned}$$

$$(2) \text{Cov}(Y, X) = \text{Cov}(X, Y)$$

$$\begin{aligned} (3) \text{Cov}(cX, Y) &= E[cXY] - E[cX]E[Y] \\ &= cE[XY] - cE[X]E[Y] \\ &= c\text{Cov}(X, Y) \end{aligned}$$

$$\begin{aligned}
 \textcircled{4} \quad \text{Cov}(X, Y+Z) &= E[X(Y+Z)] - E[X]E[Y+Z] \\
 &= E[XY + XZ] - E[X](E[Y] + E[Z]) \\
 &= \underline{E[XY]} + \underline{E[XZ]} - \underline{E[X]E[Y]} - \underline{E[X]E[Z]} \\
 &= \text{Cov}(X, Y) + \text{Cov}(X, Z)
 \end{aligned}$$

Note: The covariance is a "bilinear" operator

$$\text{Find } V\left[\sum_{i=1}^n X_i\right]$$

$$= \text{Cov}\left[\sum_{i=1}^n X_i, \sum_{j=1}^n X_j\right] \quad \text{now use } \textcircled{4} \text{ and } \textcircled{2}$$

$$= \sum_{i=1}^n \text{Cov}(X_i, X_i) + \sum_{i \neq j} \text{Cov}(X_i, X_j)$$

$$= \sum_{i=1}^n V[X_i] + 2 \sum_{i < j} \text{Cov}(X_i, X_j)$$

(7)

Note: If X_i, X_j are independent $\forall i \neq j$,
 then $V[\sum_{i=1}^n X_i] = \sum_{i=1}^n V[X_i]$

Consequence: Recall $X \sim \text{Bino}(n, p)$

$$X = \sum_{i=1}^n X_i \text{ where } X_i \sim \text{Bern}(p)$$

And the X_i 's are independent

$$\therefore V[X] = \sum_{i=1}^n V[X_i] = npq$$

(8)

Hypergeometric Distribution

You have a population of N items

R of them are of a particular type.

Randomly select n items.

$X = \#$ items of that type in the sample.

$$p(x) = \frac{\binom{R}{x} \binom{N-R}{n-x}}{\binom{N}{n}} \quad x = 0, 1, \dots, \min(R, n)$$

Let $X_i = \begin{cases} 1 & \text{If the } i^{\text{th}} \text{ item is of that type} \\ 0 & \text{otherwise} \end{cases}$ (9)

$$P[X_i = 1] = \frac{R}{N} \quad \text{Call this } p$$

True $\forall i$ because the X_i 's are "interchangeable"

$$\begin{aligned} \therefore E[X] &= E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] = np \\ &= n \frac{R}{N} \end{aligned}$$

$$X_i \sim \text{Bern}(p)$$

$$V[X_i] = pq$$

$$V[X] = V\left[\sum_{i=1}^n X_i\right] = \underbrace{\sum_{i=1}^n V[X_i]}_{npq} + 2 \sum_{i < j} \text{Cov}(X_i, X_j)$$

$$\begin{aligned} \text{Cov}(X_i, X_j) &= E[X_i X_j] - E[X_i] E[X_j] \\ &= E[X_i X_j] - p^2 \end{aligned}$$

$$\begin{aligned}
 E[X_i X_j] &= 0 \cdot P[X_i X_j = 0] + 1 \cdot P[X_i X_j = 1] \quad (11) \\
 &= P[X_i X_j = 1] \\
 &= P[X_i = 1 \text{ and } X_j = 1] \\
 &= P[X_i = 1] \cdot P[X_j = 1 | X_i = 1] \\
 &= p \cdot \frac{R-1}{N-1}
 \end{aligned}$$

$$\text{So } \text{Cov}(X_i, X_j) = p \frac{R-1}{N-1} - p^2$$

$$\begin{aligned}
 \text{And } V[X] &= npq + 2 \binom{n}{2} \left(p \frac{R-1}{N-1} - p^2 \right) \quad (12) \\
 &= npq + 2 \frac{n!}{2!(n-2)!} p \left(\frac{R-1}{N-1} - p \right) \\
 &= npq + n(n-1)p \left(\frac{R-1}{N-1} - \frac{R}{N} \right) \\
 &\quad \underbrace{\frac{NR - N - NR + R}{N(N-1)}}_{\frac{(-1 + \frac{R}{N})}{N-1}} \\
 &= npq + n(n-1)p \frac{(-1 + \frac{R}{N})}{N-1}
 \end{aligned}$$

$$= npq - \frac{n(n-1)pq}{N-1}$$

(13)

$$= npq \left[1 - \frac{(n-1)}{N-1} \right]$$

$$\frac{N-1-n+1}{N-1}$$

$$V[X] = npq \left(\frac{N-n}{N-1} \right) = n \frac{R}{N} \left(1 - \frac{R}{N} \right) \left(\frac{N-n}{N-1} \right)$$

f.p.c. = finite population correction

Defn: Suppose X_1, \dots, X_n are i.i.d.

(14)

independent, identically distributed

with mean μ and variance σ^2 .

Let $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$. Then \bar{X} is called

the sample mean.

Properties: ① $E[\bar{X}] = E\left[\frac{1}{n} \sum_{i=1}^n X_i\right]$

$$= \frac{1}{n} \sum_{i=1}^n E[X_i] = \frac{1}{n} n\mu = \mu$$

(15)

$$\begin{aligned}
 \textcircled{2} \quad V[\bar{X}] &= V\left[\frac{1}{n} \sum_{i=1}^n X_i\right] \\
 &= \frac{1}{n^2} \left[\sum_{i=1}^n V[X_i] + 0 \right] \\
 &= \frac{1}{n^2} n \sigma^2 = \frac{\sigma^2}{n}
 \end{aligned}$$

$$\begin{aligned}
 \textcircled{3} \quad \text{Cov}(\bar{X}, X_i - \bar{X}) &= \text{Cov}(\bar{X}, X_i) - \underbrace{\text{Cov}(\bar{X}, \bar{X})}_{V[\bar{X}]} \\
 &= \text{Cov}\left(\frac{1}{n} \sum_{i=1}^n X_i, X_i\right) - \frac{\sigma^2}{n}
 \end{aligned}$$

(16)

$$\begin{aligned}
 &= \frac{1}{n} \sum_{j=1}^n \text{Cov}(X_j, X_i) - \frac{\sigma^2}{n} \\
 &= \frac{1}{n} \underbrace{\text{Cov}(X_i, X_i)}_{\substack{V[X_i] \\ \sigma^2}} - \frac{\sigma^2}{n} = 0
 \end{aligned}$$

$$\text{So } \text{Cov}(\bar{X}, X_i - \bar{X}) = 0 \quad \forall i$$

(17)

Tuesday: Midterm exam

1-page of notes (front and back)

Calculator

Covers everything from the beginning thru

Tues Feb 7

53. If X is uniform over $(0, 1)$, calculate $E[X^n]$ and $\text{Var}(X^n)$.

55. Suppose that the joint probability mass function of X and Y is

$$P(X = i, Y = j) = \binom{j}{i} e^{-2\lambda} \lambda^j / j!, \quad 0 \leq i \leq j$$

- (a) Find the probability mass function of Y .
- (b) Find the probability mass function of X .
- (c) Find the probability mass function of $Y - X$.

61. Let X and W be the working and subsequent repair times of a certain machine. Let $Y = X + W$ and suppose that the joint probability density of X and Y is

$$f_{X,Y}(x, y) = \lambda^2 e^{-\lambda y}, \quad 0 < x < y < \infty$$

- (a) Find the density of X .
- (b) Find the density of Y .
- (c) Find the joint density of X and W .
- (d) Find the density of W .

76. Let X and Y be independent random variables with means μ_x and μ_y and variances σ_x^2 and σ_y^2 . Show that

$$\text{Var}(XY) = \sigma_x^2 \sigma_y^2 + \mu_y^2 \sigma_x^2 + \mu_x^2 \sigma_y^2$$

79. With $K(t) = \log(E[e^{tX}])$, show that

$$K'(0) = E[X], \quad K''(0) = \text{Var}(X)$$