

Renewal processes

Stat 528
5-23-17

Recall: $\{X_1, X_2, \dots\}$ are the interarrival times

①

$\{N(t), t \geq 0\}$ is the renewal process

$F(x)$ is the cdf of each of the X_i 's

$$S_n = \sum_{i=1}^n X_i$$

$$N(t) \geq n \iff S_n \leq t$$

$$\begin{aligned} P[N(t) = n] &= P[S_n \leq t] - P[S_{n+1} \leq t] \\ &= F_n(t) - F_{n+1}(t) \end{aligned} \quad *$$

Example $P[X_n = i] = p(1-p)^{i-1}$ i is an integer ≥ 1

②

$S_1 = X_1$ = time until the 1st success

$S_n = \sum_{i=1}^n X_i$ = time until the n^{th} success

so S_n has a negative binomial (Pascal) distribution

$$P[S_n = k] = \binom{k-1}{n-1} p^n (1-p)^{k-n} \quad k \geq n$$

Using *,

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$$\begin{aligned} P[N(t)=n] &= F_n(t) - F_{n+1}(t) \\ &= P[S_n \leq t] - P[S_{n+1} \leq t] \\ &= \sum_{k=n}^{\lfloor t \rfloor} P[S_n = k] - \sum_{k=n+1}^{\lfloor t \rfloor} P[S_{n+1} = k] \\ &= \sum_{k=n}^{\lfloor t \rfloor} \binom{k-1}{n-1} p^n q^{k-n} - \sum_{k=n+1}^{\lfloor t \rfloor} \binom{k-1}{n} p^{n+1} q^{k-n-1} \end{aligned}$$

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Also,

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$$P[N(t)=n] = \binom{\lfloor t \rfloor}{n} p^n q^{\lfloor t \rfloor - n}$$

Another identity:

$$P[N(t)=n] = \int_0^{\infty} P[N(t)=n | S_n=y] f_{S_n}(y) dy$$

If the n^{th} event occurs at time y and $y > t$,
then $N(t) < n$, so the integrand is 0.

$$P[N(t)=n] = \int_0^t P[N(t)=n | S_n=y] f_{S_n}(y) dy \quad (5)$$

If the n^{th} event occurs at time y and $y \leq t$ then $N(t)=n$ only if no more events occurred between times y and t , i.e. the $(n+1)^{\text{th}}$ interarrival time is greater than $t-y$

$$P[N(t)=n] = \int_0^t \underbrace{P[X_{n+1} > t-y]}_{1-F(t-y)} f_{S_n}(y) dy$$

$$P[N(t)=n] = \int_0^t \bar{F}(t-y) f_{S_n}(y) dy \quad ** \quad (6)$$

Example Suppose $F(x) = 1 - e^{-\lambda x}$
 $\bar{F}(x) = 1 - F(x) = e^{-\lambda x}$
 $f(x) = \frac{d}{dx} F(x) = \lambda e^{-\lambda x} \sim \text{Exp}(\lambda)$
 $\sum_{i=1}^n S_n = \sum_{i=1}^n X_i \sim \text{Gamma}(n, \lambda)$

$$\text{Using } **, P[N(t)=n] = \int_0^t e^{-\lambda(t-y)} \frac{\lambda e^{-\lambda y} (\lambda y)^{n-1}}{(n-1)!} dy$$

$$= \frac{e^{-\lambda t} \lambda^n}{(n-1)!} \underbrace{\int_0^t y^{n-1} dy}_{\frac{y^n}{n} \Big|_0^t = \frac{t^n}{n}} \quad (7)$$

$$= \frac{e^{-\lambda t} (\lambda t)^n}{n!} \sim \text{Poisson}(\lambda t)$$

Definition $m(t) = E[N(t)]$ is the
mean-value function or the renewal function

Yet another identity:

Suppose that the r.v. Y is non-negative, integer-valued.

$$\begin{aligned} \text{Then } \sum_{n=1}^{\infty} P[Y=n] &= \sum_{n=1}^{\infty} \sum_{k=n}^{\infty} P[Y=k] \\ &= \sum_{k=1}^{\infty} \sum_{n=1}^k P[Y=k] = \sum_{k=1}^{\infty} k P[Y=k] \\ &= \sum_{k=0}^{\infty} k P[Y=k] = E[Y] \end{aligned}$$

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Use this:

$$\begin{aligned}
 m(t) &= E[N(t)] = \sum_{n=1}^{\infty} P[N(t) \geq n] \\
 &= \sum_{n=1}^{\infty} P[S_n \leq t] \\
 m(t) &= \sum_{n=1}^{\infty} F_n(t)
 \end{aligned}$$

Assume that F is continuous

$$m(t) = E[N(t)] = \int_0^{\infty} E[N(t) | X_1 = x] f(x) dx$$

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Also, when $x > t$

$$E[N(t) | X_1 = x] = 0$$

And when $x < t$, $E[N(t) | X_1 = x] = 1 + E[N(t-x)]$

$$\begin{aligned}
 \text{So } m(t) &= \int_0^t (1 + E[N(t-x)]) f(x) dx \\
 &= \int_0^t f(x) dx + \int_0^t m(t-x) f(x) dx
 \end{aligned}$$

$m(t) = F(t) + \int_0^t m(t-x) f(x) dx$	Renewal Equation
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Example Suppose that the interarrival times
are $U_{int}(0,1)$

Let $0 \leq t \leq 1$

$$m(t) = t + \int_0^t m(t-x) \cdot 1 dx \quad \text{Let } y = t-x$$

$$= t + \int_t^0 m(y) (-dy)$$

$$m(t) = t + \int_0^t m(y) dy \quad \text{Differentiate w.r.t } t$$

$$m'(t) = 1 + m(t)$$

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$$\text{Let } h(t) = 1 + m(t)$$

$$h'(t) = m'(t)$$

$$h'(t) = h(t) \quad h(t) = Ke^t$$

$$m(t) = Ke^t - 1$$

$$0 = m(0) = K - 1 \quad \therefore K = 1$$

$$\text{And } m(t) = e^t - 1$$