

Brownian Motion

Start 528

6-6-17

A random walk is a Markov chain

①

where $P_{i,i+1} = p$, $P_{i,i-1} = 1-p$, i is any integer

A symmetric random walk is a random walk with $p = .5$

Suppose that, at each time Δt we take a step of size Δx . Let $X(t)$ be our position at time t .

Let $X_i = \begin{cases} 1 & \text{if the } i^{\text{th}} \text{ step was to the right} \\ -1 & \text{if the } i^{\text{th}} \text{ step was to the left} \end{cases}$ ②

So X_1, X_2, \dots are iid random variables

with $P(X_i = 1) = P(X_i = -1) = \frac{1}{2}$

$$E[X_i] = 0 \quad V[X_i] = E[X_i^2] = 1$$

$$X(t) = X_1 \Delta x + X_2 \Delta x + \dots + X_{\lfloor t/\Delta t \rfloor} \Delta x$$

$$E[X(t)] = 0$$

$$\begin{aligned} V[X(t)] &= (\Delta x)^2 V[X_1] + \dots + (\Delta x)^2 V[X_{t/\Delta t}] \quad \textcircled{3} \\ &= (\Delta x)^2 [t/\Delta t] \end{aligned}$$

Now, let both Δx and Δt go to zero,

so that $\frac{(\Delta x)^2}{\Delta t}$ remains constant $= \sigma^2$

Then $V[X(t)] \rightarrow \sigma^2 t$

Notes: - Since $X(t)$ was the sum of iid random variables, the central limit theorem says that $X(t)$ will have a limiting normal distribution,
 i.e. $X(t) \xrightarrow{D} N(0, \sigma^2 t)$ \textcircled{4}

- Since changes in the value of $X(t)$ in nonoverlapping time intervals are independent,

$\{X(t), t \geq 0\}$ has independent increments,

c.e. $X(t_1), X(t_2) - X(t_1), X(t_3) - X(t_2),$ ⑤
 $\dots, X(t_n) - X(t_{n-1})$ are independent
for $0 < t_1 < t_2 < \dots < t_n$

- $X(t)$ has stationary increments,
c.e. $X(t+s) - X(t)$ depends only on s
(not t)

Defn: A stochastic process $\{X(t), t \geq 0\}$ ⑥
is a Brownian Motion process if

- (i) $X(0) = 0$
- (ii) $\{X(t), t \geq 0\}$ has stationary and independent increments
- (iii) $\forall t, X(t) \sim N(0, \sigma^2 t)$

If $\sigma = 1$, then $\{X(t), t \geq 0\}$ is a
standard Brownian Motion process.

Note: If $X(t)$ is a Brownian motion process, (7)
then $\frac{X(t)}{\sigma}$ is a standard ...

Since $X(t) \sim N(0, t)$, (Assume $\sigma=1$
without loss of
generality)
$$f_t(x) = \frac{1}{\sqrt{2\pi t}} e^{-\frac{1}{2} \frac{x^2}{t}}$$

Find the joint density of $X(t_1), X(t_2), \dots, X(t_n)$
for $0 < t_1 < t_2 < \dots < t_n$

The event $\begin{cases} X(t_1) = \mu_1, \\ X(t_2) = \mu_2 \\ \vdots \\ X(t_n) = \mu_n \end{cases}$ (8)

is equivalent to the event $\begin{cases} X(t_1) = \mu_1, \\ X(t_2) - X(t_1) = \mu_2 - \mu_1, \\ \vdots \\ X(t_n) - X(t_{n-1}) = \mu_n - \mu_{n-1} \end{cases}$

$$X(t_k) - X(t_{k-1}) \sim N(0, t_k - t_{k-1})$$

Now

(9)

$$f(x_1, \dots, x_n) = f_{t_1}(x_1) f_{t_2-t_1}(x_2-x_1) \dots f_{t_n-t_{n-1}}(x_n-x_{n-1})$$

$$= \frac{1}{\sqrt{2\pi t_1}} e^{-\frac{1}{2} \frac{x_1^2}{t_1}} \prod_{k=2}^n \frac{1}{\sqrt{2\pi(t_k-t_{k-1})}} e^{-\frac{1}{2} \frac{(x_k-x_{k-1})^2}{t_k-t_{k-1}}}$$

Use this to find the conditional distribution of $X(s)$, given $X(t) = B$, where $s < t$

$$f_{s|t}(x|B) = \frac{f_{s,t}(x, B)}{f_t(B)}$$

(10)

$$= \frac{\frac{1}{\sqrt{2\pi s}} e^{-\frac{1}{2} \frac{x^2}{s}}}{\frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{1}{2} \frac{(B-x)^2}{t-s}}}$$

$$= \underbrace{k}_{\substack{\uparrow \\ \text{free of } x}} e^{-\frac{1}{2} \left[\frac{x^2}{s} + \frac{(B-x)^2}{t-s} \right]}$$

$$\begin{aligned}
&= K e^{-\frac{1}{2} \left[\frac{(t-s)x^2 + s(B^2 - 2Bx + x^2)}{s(t-s)} \right]} \quad (11) \\
&= K' e^{-\frac{1}{2} \left[\frac{tx^2 + x(-2sB)}{s(t-s)} \right]} \\
&= K'' e^{-\frac{1}{2} \left[\frac{t \left[x^2 - \frac{2sB}{t}x + \frac{s^2B^2}{t^2} \right]}{s(t-s)} \right]} \\
&= K'' e^{-\frac{1}{2} \frac{\left(x - \frac{s}{t}B\right)^2}{\frac{s}{t}(t-s)}}
\end{aligned}$$

$$\text{So } X(s) | X(t)=B \sim N\left(\frac{s}{t}B, \frac{s}{t}(t-s)\right)$$

Example: Bike race with 2 competitors

Let $Y(t)$ = Amount of time (in seconds)

that the inside position is ahead when
100t% of the race is completed.

Assume $\{Y(t), 0 \leq t \leq 1\}$ is a Brownian
motion process with variance parameter σ^2 .

If the inside racer is ahead by σ seconds
At the midpoint of the race, find the prob
that he/she wins.

(13)

$$\begin{aligned} & P[Y(1) > 0 \mid Y(\tfrac{1}{2}) = \sigma] \\ &= P[Y(1) - Y(\tfrac{1}{2}) > -\sigma \mid Y(\tfrac{1}{2}) = \sigma] \\ &= P[Y(1) - Y(\tfrac{1}{2}) > -\sigma] \quad \text{by independent increments} \\ &= P[Y(\tfrac{1}{2}) > -\sigma] \quad \text{by stationary increments} \end{aligned}$$

$$\text{Also } Y(\tfrac{1}{2}) \sim N(0, \sigma^2 \tfrac{1}{2})$$

(14)

$$\text{So } \frac{Y(\tfrac{1}{2}) - 0}{\sigma/\sqrt{2}} \sim N(0, 1)$$

$$\begin{aligned} P[Y(\tfrac{1}{2}) > -\sigma] &= P[Z > \frac{-\sigma}{\sigma/\sqrt{2}}] \\ &= P[Z > -\sqrt{2}] = .9214 \end{aligned}$$

If the inside racer wins by σ seconds,
 And the probability that he/she was
 ahead at the midpoint.

(15)

$$P[Y(\frac{1}{2}) > 0 \mid Y(1) = \sigma]$$

Need the conditional distribution of $Y(s)$,
 given $Y(t) = C$, for $s < t$

Let $X(t) = \frac{Y(t)}{\sigma}$. Then $X(t)$ is a stan.
 B. mo. process

We showed the cond. dist. of $X(s)$, given $X(t) = B$ (16)

$$\text{is } N\left(\frac{s}{t}B, \frac{s}{t}(t-s)\right)$$

\uparrow
 $\frac{C}{\sigma}$

$$Y(s) = \sigma X(s)$$

$$\text{so } Y(s) \mid Y(t) = C \sim N\left(\frac{s}{t}C, \frac{s}{t}(t-s)\sigma^2\right)$$

$$P[Y(\frac{1}{2}) > 0 \mid Y(1) = \sigma] = P[W > 0] \text{ where}$$

$$W \sim N\left(\frac{1}{2}\sigma, \frac{1}{2}(1-\frac{1}{2})\sigma^2\right)$$

(17)

$$W \sim N\left(\frac{\sigma}{2}, \frac{\sigma^2}{4}\right)$$

$$\begin{aligned} P[W > 0] &= P\left[Z > \frac{0 - \frac{\sigma}{2}}{\frac{\sigma}{2}}\right] \\ &= P[Z > -1] = .8413 \end{aligned}$$