

Rank-based ANOVA

Stat 525
3-6-18
①

Let's revisit the 1-way ANOVA

$$y_{ij} = \mu + \tau_i + \varepsilon_{ij} \quad \begin{array}{l} i = 1, \dots, a \\ j = 1, \dots, n \end{array} \quad N = na$$

Usual assumptions: $\varepsilon_{ij} \sim \text{iid } N(0, \sigma^2)$

If the assumptions are not met, and n is small,
then an alternative method is recommended.

Apply the rank transformation.

②

Rank all N of the data values
from 1 (smallest) to N (largest)

- If there are ties, average the ranks for
the items involved in the tie.

Compute F :

$$SS_{\text{Tot}} = \sum_{i=1}^a \sum_{j=1}^n (y_{ij} - \bar{y}_{..})^2$$

After ranking,
$$SS_{\text{Tot}} = \sum_{i=1}^a \sum_{j=1}^n (R_{ij} - \bar{R}_{..})^2$$

Sum of integers from 1 to N is $\frac{N(N+1)}{2}$

③

Sum of squares from 1 to N is $\frac{N(N+1)(2N+1)}{6}$

$$\bar{R}_{..} = \frac{1}{N} \sum_{i=1}^a \sum_{j=1}^b R_{ij} = \frac{1}{N} \frac{N(N+1)}{2} = \frac{N+1}{2}$$

$$\begin{aligned} SS_{TOT} &= \sum_{i=1}^a \sum_{j=1}^b (R_{ij} - \bar{R}_{..})^2 = \sum_i \sum_j R_{ij}^2 - N \bar{R}_{..}^2 \\ &= \frac{N(N+1)(2N+1)}{6} - N \left(\frac{N+1}{2} \right)^2 = \frac{N(N^2-1)}{12} \end{aligned}$$

This implies that $SS_{TRT} + SS_{ERR} = \text{Constant}$,

④

So $MS_{TRT} \nmid MS_{ERR}$ are no longer independent.

$$\begin{aligned} \text{Compute } SS_{TRT} &= \sum_{i=1}^a \sum_{j=1}^b (\bar{y}_{i.} - \bar{y}_{..})^2 \\ &= n \left(\sum_{i=1}^a \bar{y}_{i.}^2 - a \bar{y}_{..}^2 \right) \end{aligned}$$

After ranking,

$$SS_{TRT} = n \left(\sum_{i=1}^a \bar{R}_{i.}^2 - a \bar{R}_{..}^2 \right)$$

$$= n \sum_{i=1}^a \bar{R}_{i.}^2 - N \left(\frac{N+1}{2} \right)^2 \quad (5)$$

$$= n \sum_{i=1}^a \left(\frac{R_{i.}}{n} \right)^2 - \frac{N(N+1)^2}{4}$$

$$SS_{TRT} = \frac{1}{n} \sum_{i=1}^a R_{i.}^2 - \frac{N(N+1)^2}{4}$$

In the usual F test, we have that

$$\text{Under } H_0, \quad \frac{SS_{TRT}}{\sigma^2} \sim \chi^2_{a-1}$$

But after ranking, what is σ^2 ? (6)

Before, it was $V(\varepsilon_{ij}) = V(y_{ij})$

$$\text{Now, } \sigma^2 = V(R_{ij})$$

$$= V(\text{discrete uniform distribution})$$

$$= \frac{N^2 - 1}{12}$$

$$= \frac{1}{n} SS_{TOT}$$

$$\text{Now } \frac{SS_{\text{TRT}}}{\sigma^2} = \frac{\frac{1}{n} \sum_{i=1}^a R_i^2 - \frac{N(N+1)^2}{4}}{\frac{N^2-1}{12}} \quad (7)$$

$$= \frac{12}{N^2-1} \left[\frac{1}{n} \sum_{i=1}^a R_i^2 - \frac{N(N+1)^2}{4} \right]$$

$$= \frac{N}{N-1} \left[\underbrace{\frac{12}{N(N+1)} \sum_{i=1}^a \frac{R_i^2}{n}}_H - 3(N+1) \right]$$

H is the Kruskal-Wallis test statistic

H has an approximate χ^2_{a-1} distribution. (8)

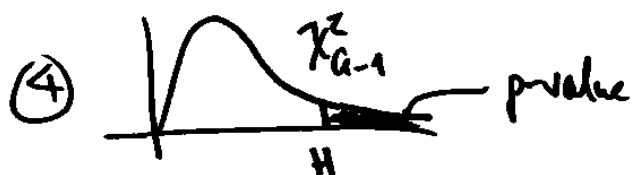
(So does $\frac{N}{N-1} H$)

Kruskal-Wallis test:

① Rank the data

② Find the sum of the ranks for each of the a levels of the factor

③ Compute H



(9)

But what is actually being tested?

$$E[g(Y)] \neq g(E[Y])$$

So we are no longer testing $H_0: \mu_1 = \mu_2 = \dots = \mu_a$

After ranking, you are actually testing the null hypothesis that the population medians are all equal.

(10)

#2 from in-class portion of mid-term

$$Y_{ijk} = \mu + \tau_i + \gamma_{ij} + \varepsilon_{ijk} \quad \begin{array}{l} i=1,2,3 \\ j=1,2,3 \\ k=1,2,3 \end{array}$$

$$\sum_{i=1}^3 \tau_i = 0 \quad \sum_{j=1}^3 \gamma_{ij} = 0$$

Write the matrix form

(11)

y_{111}
y_{112}
y_{113}
y_{121}
y_{122}
y_{123}
y_{131}
y_{132}
y_{133}
...
...
...
...
y_{332}
y_{333}

=

1	1 0	1 0 0 0 0 0
1	1 0	1 0 0 0 0 0
1	1 0	1 0 0 0 0 0
1	1 0	0 1 0 0 0 0
1	1 0	0 1 0 0 0 0
1	1 0	0 1 0 0 0 0
1	1 0	-1 -1 0 0 0 0
1	1 0	-1 -1 0 0 0 0
1	1 0	-1 -1 0 0 0 0
...
...
...
...
1	-1 -1	0 0 0 0 -1 -1
1	-1 -1	0 0 0 0 -1 -1

27x1

μ
γ_1
γ_2
γ_{11}
γ_{12}
γ_{21}
γ_{22}
γ_{31}
γ_{32}
γ_{33}

+ $\sum_{i=1}^{27} \gamma_i$

Source	df
A	2
G	6
ERR	18
Tot	26

27x9

(12)

#3 from in-class midterm

$$a = 3 \quad k = 2 \quad \lambda = 2$$

$$N = ar = bk \quad \text{and} \quad \lambda = \frac{r(k-1)}{a-1}$$

$$2 = \frac{r(2-1)}{3-1}$$

$$\therefore \underline{r = 4} \quad (b)$$

$$N = 3 \cdot 4 = b \cdot 2$$

$$\therefore \underline{b = 6} \quad (a)$$

(13)

(c)

		BLK					
		1	2	3	4	5	6
1		✓	✓		✓	✓	
2		✓		✓	✓		✓
3			✓	✓		✓	✓