

Simple linear regression (continued) Stat 564
9-28-17

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i \quad E[\varepsilon_i] = 0 \quad \forall i \quad (1)$$

$$V[\varepsilon_i] = \sigma^2 \quad \forall i$$

$\varepsilon_i, \varepsilon_j$ are independent
 $\forall i \neq j$

Find the properties of $\hat{\beta}_0$ and $\hat{\beta}_1$.

$$E[\hat{\beta}_1] = E\left[\frac{S_{xy}}{S_{xx}}\right]$$

$$= \frac{1}{S_{xx}} E[S_{xy}]$$

$$= \frac{1}{S_{xx}} E\left[\sum_{i=1}^n (x_i - \bar{x}) y_i\right]$$

$$= \frac{1}{S_{xx}} \sum_{i=1}^n (x_i - \bar{x}) E[y_i]$$

$$= \frac{1}{S_{xx}} \sum_{i=1}^n (x_i - \bar{x}) E[\beta_0 + \beta_1 x_i + \varepsilon_i]$$

(2)

$$= \frac{1}{S_{xx}} \sum_{i=1}^n (x_i - \bar{x}) (\beta_0 + \beta_1 x_i) \quad (3)$$

$$= \frac{1}{S_{xx}} \left[\underbrace{\beta_0 \sum_{i=1}^n (x_i - \bar{x})}_0 + \beta_1 \underbrace{\sum_{i=1}^n (x_i - \bar{x}) x_i}_{S_{xx}} \right]$$

$$= \frac{1}{S_{xx}} [\beta_1 S_{xx}] = \beta_1$$

$\therefore \hat{\beta}_1$ is an unbiased estimator of β_1 .

$$\text{Var}(\hat{\beta}_1) = V\left(\frac{S_{xy}}{S_{xx}}\right) \quad (4)$$

$$= \frac{1}{S_{xx}^2} V(S_{xy})$$

$$= \frac{1}{S_{xx}^2} V\left(\sum_{i=1}^n (x_i - \bar{x}) y_i\right)$$

$$= \frac{1}{S_{xx}^2} \left[\sum_{i=1}^n (x_i - \bar{x})^2 V(y_i) + \sum_{i < j} \sum_{j} (x_i - \bar{x})(x_j - \bar{x}) \cancel{\text{Cov}(y_i, y_j)} \right]$$

because $\varepsilon_i, \varepsilon_j$
are indep

$$= \frac{1}{\sum_{xx}} \sum_{i=1}^n (x_i - \bar{x})^2 V(\beta_0 + \beta_1 x_i + \varepsilon_i) \quad (5)$$

$$= \frac{1}{\sum_{xx}} \sum_{i=1}^n (x_i - \bar{x})^2 \underbrace{V(\varepsilon_i)}_{\sigma^2}$$

$$= \frac{\sigma^2}{\sum_{xx}} \sum_{xx} = \frac{\sigma^2}{\sum_{xx}}$$

$$V(\hat{\beta}_1) = \frac{\sigma^2}{\sum_{xx}} = \frac{\sigma^2}{(n-1)S_x^2}$$

Since $\hat{\beta}_1$ was unbiased
 $\lim_{n \rightarrow \infty} V(\hat{\beta}_1) = 0$,
 $\hat{\beta}_1$ is consistent

What about $\hat{\beta}_0$?

$$E[\hat{\beta}_0] = E[\bar{y} - \hat{\beta}_1 \bar{x}] = E[\bar{y}] - \bar{x} E[\hat{\beta}_1]$$

$$= E\left[\frac{1}{n} \sum_{i=1}^n y_i\right] - \bar{x} \beta_1$$

$$= \frac{1}{n} \sum_{i=1}^n E(\beta_0 + \beta_1 x_i + \varepsilon_i) - \bar{x} \beta_1$$

$$= \frac{1}{n} \sum_{i=1}^n (\beta_0 + \beta_1 x_i + 0) - \bar{x} \beta_1$$

$$= \frac{1}{n} (n \beta_0 + \beta_1 \sum x_i) - \bar{x} \beta_1 = \beta_0$$

$\therefore \hat{\beta}_0$ is
 also
 unbiased

$$V[\hat{\beta}_0] = V[\bar{y} - \hat{\beta}_1 \bar{x}] \quad (7)$$

$$= V\left[\frac{1}{n} \sum_{i=1}^n y_i - \bar{x} \frac{1}{S_{xx}} \sum_{i=1}^n (x_i - \bar{x}) y_i\right]$$

$$= V\left[\sum_{i=1}^n \underbrace{\left(\frac{1}{n} - \frac{\bar{x}(x_i - \bar{x})}{S_{xx}}\right)}_{k_i} y_i\right]$$

$$= \sum_{i=1}^n k_i^2 \underbrace{V[y_i]}_{\sigma^2} + 0$$

$$= \sigma^2 \sum_{i=1}^n \left(\frac{1}{n} - \frac{\bar{x}(x_i - \bar{x})}{S_{xx}}\right)^2$$

$$= \sigma^2 \sum_{i=1}^n \left(\frac{1}{n^2} + \frac{\bar{x}^2 (x_i - \bar{x})^2}{S_{xx}^2} - \frac{2\bar{x}(x_i - \bar{x})}{n S_{xx}}\right) \quad (8)$$

$$= \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}^2} \underbrace{\sum_{i=1}^n (x_i - \bar{x})^2}_{S_{xx}} - \frac{2\bar{x}}{n S_{xx}} \underbrace{\sum_{i=1}^n (x_i - \bar{x})}_0 \right]$$

$$V(\hat{\beta}_0) = \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}} \right]$$

$$= \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{(n-1)S_x^2} \right]$$

$\therefore \hat{\beta}_0$ is also
a consistent
estimator

We need to find an estimator of σ^2 .

(9)

$$\begin{aligned} SSE &= \sum_{i=1}^n (y_i - \hat{y}_i)^2 \\ &= \sum_{i=1}^n (y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i))^2 \\ &= \sum_{i=1}^n (y_i - (\bar{y} - \hat{\beta}_1 \bar{x}) - \hat{\beta}_1 x_i)^2 \\ &= \sum_{i=1}^n [(y_i - \bar{y}) - \hat{\beta}_1 (x_i - \bar{x})]^2 \end{aligned}$$

$$= \sum_{i=1}^n [(y_i - \bar{y})^2 + \hat{\beta}_1^2 (x_i - \bar{x})^2 - 2\hat{\beta}_1 (y_i - \bar{y})(x_i - \bar{x})] \quad (10)$$

$$SSE = S_{yy} + \hat{\beta}_1^2 S_{xx} - 2\hat{\beta}_1 S_{xy}$$

$$= S_{yy} + \left(\frac{S_{xy}}{S_{xx}}\right)^2 S_{xx} - 2\left(\frac{S_{xy}}{S_{xx}}\right) S_{xy}$$

$$SSE = S_{yy} - \frac{S_{xy}^2}{S_{xx}} = S_{yy} - \hat{\beta}_1^2 S_{xx}$$

⑪

$$\begin{aligned}
 E(SSE) &= E[S_{yy} - \hat{\beta}_1^2 S_{xx}] \\
 &= \underbrace{E[S_{yy}]}_{(1)} - S_{xx} \underbrace{E[\hat{\beta}_1^2]}_{(2)}
 \end{aligned}$$

$$(2): V[X] = E[X^2] - (E[X])^2 \Rightarrow E[X^2] = V[X] + (E[X])^2$$

$$\begin{aligned}
 E[\hat{\beta}_1^2] &= V[\hat{\beta}_1] + (E[\hat{\beta}_1])^2 \\
 &= \frac{\sigma^2}{S_{xx}} + \beta_1^2
 \end{aligned}$$

⑫

$$\begin{aligned}
 (1): S_{yy} &= \sum_{i=1}^n (y_i - \bar{y})^2 \\
 &= \sum_{i=1}^n (\beta_0 + \beta_1 x_i + \varepsilon_i - (\beta_0 + \beta_1 \bar{x} + \bar{\varepsilon}))^2 \\
 &= \sum_{i=1}^n [\beta_1 (x_i - \bar{x}) + (\varepsilon_i - \bar{\varepsilon})]^2 \\
 &= \sum_{i=1}^n [\beta_1^2 (x_i - \bar{x})^2 + (\varepsilon_i - \bar{\varepsilon})^2 + 2\beta_1 (x_i - \bar{x})(\varepsilon_i - \bar{\varepsilon})]
 \end{aligned}$$

$$S_{yy} = \beta_1^2 S_{xx} + \sum_{i=1}^n (e_i - \bar{e})^2 + 2\beta_1 \sum_{i=1}^n (x_i - \bar{x})(e_i - \bar{e}) \quad (13)$$

$$\begin{aligned} E[S_{yy}] &= \beta_1^2 S_{xx} + E[(n-1)\sigma_e^2] + 2\beta_1 \sum_{i=1}^n (x_i - \bar{x}) E(e_i - \bar{e}) \\ &= \beta_1^2 S_{xx} + (n-1)\sigma^2 \end{aligned}$$

$$E(SSE) = (1) - S_{xx}(2)$$

$$= \beta_1^2 S_{xx} + (n-1)\sigma^2 - S_{xx} \left(\frac{\sigma^2}{S_{xx}} + \beta_1^2 \right)$$

$$= (n-1)\sigma^2 - \sigma^2 = (n-2)\sigma^2$$

$$\text{So } E\left(\frac{SSE}{n-2}\right) = \sigma^2$$

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$$\text{Defn: } MSE = \frac{SSE}{n-2} = \hat{\sigma}^2$$

↑
"mean squared error"

The MSE is an unbiased estimator of σ^2

(15)

Summary:

$$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}}, \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}, \hat{\sigma}^2 = \text{MSE} = \frac{\text{SSE}}{n-2}$$

$$E[\hat{\beta}_1] = \beta_1, E[\hat{\beta}_0] = \beta_0, E[\hat{\sigma}^2] = \sigma^2$$

$$V[\hat{\beta}_1] = \frac{\sigma^2}{S_{xx}}, V[\hat{\beta}_0] = \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}} \right]$$

$$\hat{V}[\hat{\beta}_1] = \frac{\hat{\sigma}^2}{S_{xx}}, \hat{V}[\hat{\beta}_0] = \hat{\sigma}^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}} \right]$$

$\frac{\hat{\sigma}}{\sqrt{S_{xx}}}$ and $\hat{\sigma} \sqrt{\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}}$ are the standard errors of $\hat{\beta}_1$ and $\hat{\beta}_0$

2.4 [PANKRATZ · Forecasting with of least-squares estimat](#)
[aid="5N5AR">Table B.3](#) presents data on the gasoline mileage performance of 32 different automobiles.

- a.** Fit a simple linear regression model relating gasoline mileage y (miles per gallon) to engine displacement x_1 (cubic inches).
- b.** Construct the analysis-of-variance table and test for significance of regression.
- c.** What percent of the total variability in gasoline mileage is accounted for by the linear relationship with engine displacement?
- d.** Find a 95% CI on the mean gasoline mileage if the engine displacement is 275 in.³
- e.** Suppose that we wish to predict the gasoline mileage obtained from a car with a 275-in.³ engine. Give a point estimate of mileage. Find a 95% prediction interval on the mileage.
- f.** Compare the two intervals obtained in parts d and e. Explain the difference between them. Which one is wider, and why?

2.7 The purity of oxygen produced by a fractional distillation process is thought to be related to the percentage of hydrocarbons in the

main condensor of the processing unit. Twenty samples are shown below.

- a.** Fit a simple linear regression model to the data.
- b.** Test the hypothesis $H_0: \beta_1 = 0$.
- c.** Calculate R^2 .
- d.** Find a 95% CI on the slope.
- e.** Find a 95% CI on the mean purity when the hydrocarbon percentage is 1.00.

Purity (%)	Hydrocarbon (%)	Purity (%)	Hydrocarbon (%)
86.91	1.02	96.73	1.46
89.85	1.11	99.42	1.55
90.28	1.43	98.66	1.55
86.34	1.11	96.07	1.55
92.58	1.01	93.65	1.40
87.33	0.95	87.31	1.15
86.29	1.11	95.00	1.01
91.86	0.87	96.85	0.99
95.61	1.43	85.20	0.95
89.86	1.02	90.56	0.98