

Additional assumption needed
for confidence intervals & hypothesis
tests:

Stat 524
10-3-17
①

$\varepsilon_1, \varepsilon_2, \dots, \varepsilon_n$ are independent, (c.i.d.)
identically distributed $N(0, \sigma^2)$ random variables.

- Facts:
- ① $\hat{\beta}_1$ is normally distributed
 - ② $\frac{(n-2)\hat{\sigma}^2}{\sigma^2}$ has a χ^2 distribution with $n-2$ df
 - ③ $\hat{\beta}_1$ and $\hat{\sigma}^2$ are independent

Use these facts to get:

②

$$\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}} \sim N(0, 1)$$

Recall: If $Z \sim N(0, 1)$ and $W \sim \chi^2_v$
and Z, W are independent, then

$$\frac{Z}{\sqrt{\frac{W}{v}}} \sim t_v$$

$$\frac{\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}}}{\sqrt{\frac{(n-2)\hat{\sigma}^2}{\hat{\sigma}^2/(n-2)}}} \sim t_{n-2}$$

(3)

$$\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\hat{\sigma}^2 / S_{xx}}} \sim t_{n-2}$$

Hypothesis test

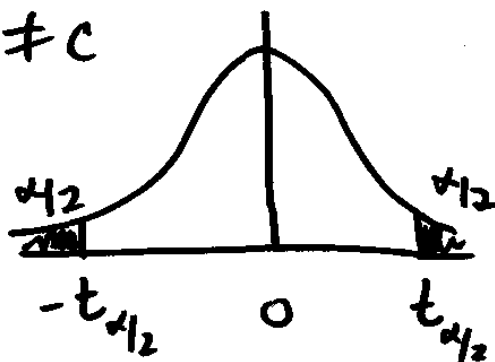
$$H_0: \beta_1 = c$$

$$H_1: \beta_1 \neq c$$

(4)

Test statistic =

$$\frac{\hat{\beta}_1 - c}{\sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}}$$



And reject H_0 if the test stat. exceeds $t_{\alpha/2}$ in absolute value

Confidence interval for β_1 is

$$\hat{\beta}_1 \pm t_{\alpha/2} \sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}$$

⑤

Observe $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$

and fit the least squares line

At a new x value, called x_0 , observe
 k new values of y .

Let \bar{y}_0 denote the average of these new y -values.

Consider the random variable $\hat{y}_0 - \bar{y}_0$ ⑥

$$E(\hat{y}_0 - \bar{y}_0) = E(\hat{\beta}_0 + \hat{\beta}_1 x_0 - (\beta_0 + \beta_1 x_0 + \bar{\epsilon}_k))$$

$$= \beta_0 + \beta_1 x_0 - \beta_0 - \beta_1 x_0 - 0$$

$$= 0$$

$$V(\hat{y}_0 - \bar{y}_0) = V(\hat{\beta}_0 + \hat{\beta}_1 x_0 - \beta_0 - \beta_1 x_0 - \bar{\epsilon}_k)$$

$$= V(\hat{\beta}_0 + \hat{\beta}_1 x_0 - \bar{\epsilon}_k)$$

$$= V(\bar{y} - \hat{\beta}_1 \bar{x} + \hat{\beta}_1 x_0 - \bar{\epsilon}_k)$$

$$= V(\bar{y} - \hat{\beta}_1(\bar{x} - \mu_0) - \bar{\varepsilon}_k) \quad (7)$$

$$= V(\bar{y} - \hat{\beta}_1(\bar{x} - \mu_0)) + V(\bar{\varepsilon}_k)$$

$$+ 2\text{Cov}$$

$$= V\left(\frac{1}{n} \sum_{i=1}^n y_i - \frac{S_{xy}}{S_{xx}} (\bar{x} - \mu_0)\right) + \frac{\sigma^2}{k}$$

$$= V\left(\frac{1}{n} \sum_{i=1}^n y_i - \left[\sum_{i=1}^n (\mu_i - \bar{x}) y_i\right] \frac{\bar{x} - \mu_0}{S_{xx}}\right) + \frac{\sigma^2}{k}$$

$$= V\left[\sum_{i=1}^n \left(\frac{1}{n} - \frac{\bar{x} - \mu_0}{S_{xx}} (\mu_i - \bar{x})\right) y_i\right] + \frac{\sigma^2}{k}$$

$$= \sum_{i=1}^n \left(\frac{1}{n} - \frac{\bar{x} - \mu_0}{S_{xx}} (\mu_i - \bar{x})\right)^2 \sigma^2 + \frac{\sigma^2}{k} \quad (8)$$

$$= \sigma^2 \left[\frac{n}{n^2} + \frac{(\bar{x} - \mu_0)^2}{S_{xx}^2} \sum_{i=1}^n (\mu_i - \bar{x})^2 - \frac{2}{n} \frac{\bar{x} - \mu_0}{S_{xx}} \sum_{i=1}^n (\mu_i - \bar{x}) + \frac{1}{k} \right]$$

$$= \sigma^2 \left[\frac{1}{n} + \frac{(\bar{x} - \mu_0)^2}{S_{xx}} + \frac{1}{k} \right]$$

$$= V(\hat{y}_0 - \bar{y}_0)$$

Consequences:

(9)

- ① Prediction interval for a single new y-value at $x = x_0$ ($k=1$)

$$\hat{y}_0 \pm t_{\alpha/2} \hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}}}$$

- ② Confidence interval for the mean of all possible y-values at $x = x_0$ ($k=\infty$)

$$\hat{y}_0 \pm t_{\alpha/2} \hat{\sigma} \sqrt{\frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}}}$$

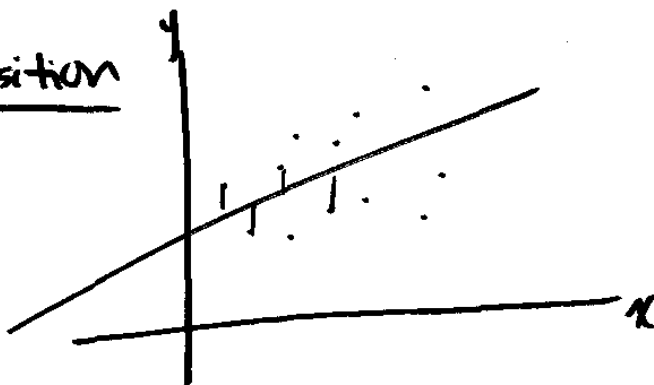
- ③ Confidence interval for β_0

(10)

($x_0=0, k=\infty$)

$$\hat{\beta}_0 \pm t_{\alpha/2} \hat{\sigma} \sqrt{\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}}$$

Variance decomposition



(11)

$$\begin{aligned}
 S_{yy} &= \sum_{i=1}^n (y_i - \bar{y})^2 \\
 &= \sum_{i=1}^n (y_i - \hat{y}_i + \hat{y}_i - \bar{y})^2 \\
 &= \sum_{i=1}^n (y_i - \hat{y}_i)^2 + \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 \\
 &\quad + 2 \sum_{i=1}^n (y_i - \hat{y}_i)(\hat{y}_i - \bar{y})
 \end{aligned}$$

①: SSE, defined last time

(12)

$$\begin{aligned}
 \textcircled{2} \quad & \sum_{i=1}^n (\hat{\beta}_0 + \hat{\beta}_1 x_i - \bar{y})^2 \\
 &= \sum_{i=1}^n (\cancel{\bar{y}} - \hat{\beta}_1 \bar{x} + \hat{\beta}_1 x_i - \cancel{\bar{y}})^2 \\
 &= \sum_{i=1}^n [\hat{\beta}_1 (x_i - \bar{x})]^2 = \hat{\beta}_1^2 \underbrace{\sum_{i=1}^n (x_i - \bar{x})^2}_{S_{xx}} \\
 &= \left[\frac{S_{xy}}{S_{xx}} \right]^2 S_{xx} = \frac{S_{xy}^2}{S_{xx}}
 \end{aligned}$$

Recall from last time: $SSE = S_{yy} - \frac{S_{xy}^2}{S_{xx}}$ (13)

$$\Rightarrow S_{yy} = SSE + \frac{S_{xy}^2}{S_{xx}}$$

$$\Rightarrow (3) = 0$$

$$\therefore S_{yy} = \sum_{i=1}^n (y_i - \hat{y}_i)^2 + \sum_{i=1}^n (\hat{y}_i - \bar{y})^2$$

$$SST = \underset{\substack{\uparrow \\ \text{total}}}{SSE} + \underset{\substack{\uparrow \\ \text{error}}}{SSR} + \underset{\substack{\uparrow \\ \text{regression}}}{SSR}$$

Definition:

$$R^2 = \frac{SSR}{SST} = \text{proportion of the variation in } y \text{ that is explained by } x. \quad (14)$$

$$0 \leq R^2 \leq 1$$

\uparrow
 x , as a linear predictor, has no effect on y

\uparrow
 perfect fit of line

$R^2 = 0:$

