

Chapter 6

Stat 562  
2-14-19

let  $X_1, \dots, X_n \sim \text{iid } f(x|\theta)$

①

Then the joint pdf of  $X_1, \dots, X_n$  is

$$f(\vec{x}|\theta) = \prod_{i=1}^n f(x_i|\theta)$$

let  $T(\vec{X})$  be a function of  $X_1, \dots, X_n$ ,  
but free of  $\theta$ .

let  $g(t|\theta)$  be the pdf of the rv.  $T(\vec{X})$  ②

Then the conditional distribution of  $\vec{X}$ , given  $T$

is 
$$\frac{\prod_{i=1}^n f(x_i|\theta)}{g(t|\theta)}$$

Defn: If this ratio  is free of  $\theta$ ,

then  $T$  is a sufficient statistic for  $\theta$ .

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Example:  $X_1, \dots, X_n \sim \text{iid } N(\theta, \sigma^2)$   
 $\uparrow$  known

$$\text{let } T = \sum_{i=1}^n X_i$$

Show that  $T$  is sufficient for  $\theta$ .

$$T \sim N(n\theta, n\sigma^2)$$

$$\frac{\prod_{i=1}^n f(x_i|\theta)}{g(t|\theta)} = \frac{\prod_{i=1}^n \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x_i - \theta}{\sigma}\right)^2}}{\frac{1}{\sigma n \sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{t - n\theta}{\sigma n}\right)^2}}$$

$$= \frac{\sigma n \sqrt{2\pi}}{\sigma^n (\sqrt{2\pi})^n} \frac{e^{-\frac{1}{2\sigma^2}(\sum x_i^2 - 2\theta \sum x_i + n\theta^2)}}{e^{-\frac{1}{2n\sigma^2}(t^2 - 2nt\theta + n^2\theta^2)}} \quad (4)$$

$\therefore$  this is free of  $\theta$ .

Example:  $X_1, \dots, X_n \sim \text{iid } f(x|\theta) = e^{-(x-\theta)}$   
 $x > \theta$

$$\text{let } T = X_{(1)}$$

We know from a previous example that

$T$  will have an exponential distribution  
 with scale parameter  $n$ .

$$\frac{f(\vec{x}|\theta)}{g(t|\theta)} = \frac{\prod_{i=1}^n e^{-(x_i - \theta)}}{n e^{-n(t - \theta)}} = \frac{e^{-\sum x_i}}{n e^{-nt}} \quad (5)$$

$\therefore T = X_{(1)}$  is sufficient for  $\theta$ .

Theorem (Neyman):  $T$  is a sufficient statistic for  $\theta$  iff

$$\prod_{i=1}^n f(x_i|\theta) = k(t|\theta) \underbrace{h(\vec{x})}_{\text{free of } \theta}$$

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Proof: ( $\Rightarrow$ ) If  $T$  is sufficient,

$$\text{then } \frac{f(\vec{x}|\theta)}{g(t|\theta)} = \underbrace{h(\vec{x})}_{\text{free of } \theta}$$

$$\hookrightarrow f(\vec{x}|\theta) = g(t|\theta) h(\vec{x})$$

$$(\Leftarrow) \text{ Assume } \prod_{i=1}^n f(x_i|\theta) = k(t|\theta) h(\vec{x})$$

$$\text{let } T_1 = T$$

⑦

$$\begin{array}{l} \text{Create } T_2 = u_2(x_1, \dots, x_n) \\ \vdots \\ T_n = u_n(x_1, \dots, x_n) \end{array} \left. \vphantom{\begin{array}{l} T_2 \\ \vdots \\ T_n \end{array}} \right\} \begin{array}{l} \text{So that the} \\ \text{system is} \\ \text{invertible} \end{array}$$

$$\begin{array}{l} \text{So } x_1 = v_1(T_1, \dots, T_n) \\ x_2 = v_2(T_1, \dots, T_n) \\ \vdots \\ x_n = v_n(T_1, \dots, T_n) \end{array} \left. \vphantom{\begin{array}{l} x_1 \\ x_2 \\ \vdots \\ x_n \end{array}} \right\} \begin{array}{l} \text{Find } J. \\ \text{It is free} \\ \text{of } \theta. \end{array}$$

⑧

$$\begin{aligned} q(t_1, \dots, t_n) &= f(x_1, \dots, x_n) |J| \\ &= k(t|\theta) \underbrace{h(\vec{x}) |J|}_{\text{free of } \theta} \end{aligned}$$

$$\begin{aligned} q(t_1) &= \underbrace{\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty}}_{n-1} k(t|\theta) h(\vec{x}) |J| dt_2 \dots dt_n \\ &= k(t_1|\theta) \underbrace{\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} h(\vec{x}) |J| dt_2 \dots dt_n}_{\substack{\text{function of } t_1, \dots, t_n \\ \text{can only be a function of } t_1}} \end{aligned}$$

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$$\sum g(t) = k(t|\theta) \underbrace{m(t)}_{\text{free of } \theta}$$

$$\text{Then } \frac{f(\vec{x}|\theta)}{g(t|\theta)} = \frac{k(t|\theta) h(\vec{x})}{\cancel{k(t|\theta)} m(t)}, \text{ free of } \theta$$

$\sum T$  is sufficient.

$$\text{Example: } X_1, \dots, X_n \sim \text{iid } f(x|\theta) = \theta x^{\theta-1} \quad 0 < x < 1 \quad (10)$$

Find a sufficient statistic for  $\theta$ .

$$\begin{aligned} f(\vec{x}|\theta) &= \prod_{i=1}^n \theta x_i^{\theta-1} = \theta^n \left( \prod_{i=1}^n x_i \right)^{\theta-1} \\ &= \frac{\theta^n \left( \prod_{i=1}^n x_i \right)^{\theta-1}}{1} \end{aligned}$$

$$\sum T = \prod_{i=1}^n x_i \text{ is sufficient for } \theta.$$

(11)

Note: the factorization is not unique

$$f(x|\theta) = \underbrace{\theta^n \left(\prod_{i=1}^n x_i\right)^\theta} \cdot \underbrace{\left(\prod_{i=1}^n x_i\right)^{-1}}$$

Defn:  $T$  is a minimal sufficient statistic for  $\theta$  if, for any other sufficient statistic  $T^*$  of  $\theta$ ,  $T$  is a function of  $T^*$ .

(12)

Example:  $X_1, \dots, X_n \sim \text{iid } N(\theta, \sigma^2)$   
 $\uparrow$  known

$T = \bar{X}$  was sufficient for  $\theta$ .

let  $T^* = (\bar{X}, S^2)$

$T$  is a function of  $T^*$ ,

but  $T^*$  is not a function of  $T$ .

So  $T^*$  is definitely not minimal

Theorem: let  $X_1, \dots, X_n \sim \text{iid } f(x|\theta)$

(13)

Suppose that a statistic  $T(\vec{X})$   
has the following property:

$$* \left\{ \begin{array}{l} \text{The ratio } \frac{f(\vec{x}|\theta)}{f(\vec{y}|\theta)} \text{ is free of } \theta \\ \text{iff } T(\vec{x}) = T(\vec{y}) \end{array} \right.$$

Then  $T$  is a minimal sufficient statistic  
for  $\theta$ .

Proof: next time

(14)

HW #5 due 2/21

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                 9  
                 10  
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**6.6** Let  $X_1, \dots, X_n$  be a random sample from a  $\text{gamma}(\alpha, \beta)$  population. Find a two-dimensional sufficient statistic for  $(\alpha, \beta)$ .

**6.9** For each of the following distributions let  $X_1, \dots, X_n$  be a random sample. Find a minimal sufficient statistic for  $\theta$ .

(a)  $f(x|\theta) = \frac{1}{\sqrt{2\pi}} e^{-(x-\theta)^2/2}, \quad -\infty < x < \infty, \quad -\infty < \theta < \infty$  (normal)

(b)  $f(x|\theta) = e^{-(x-\theta)}, \quad \theta < x < \infty, \quad -\infty < \theta < \infty$  (location exponential)

(c)  $f(x|\theta) = \frac{e^{-(x-\theta)}}{(1+e^{-(x-\theta)})^2}, \quad -\infty < x < \infty, \quad -\infty < \theta < \infty$  (logistic)

(d)  $f(x|\theta) = \frac{1}{\pi[1+(x-\theta)^2]}, \quad -\infty < x < \infty, \quad -\infty < \theta < \infty$  (Cauchy)

(e)  $f(x|\theta) = \frac{1}{2} e^{-|x-\theta|}, \quad -\infty < x < \infty, \quad -\infty < \theta < \infty$  (double exponential)

**6.10** Show that the minimal sufficient statistic for the  $\text{uniform}(\theta, \theta + 1)$ , found in Example 6.2.15, is not complete.

**6.13** Suppose  $X_1$  and  $X_2$  are iid observations from the pdf  $f(x|\alpha) = \alpha x^{\alpha-1} e^{-x^\alpha}, x > 0, \alpha > 0$ . Show that  $(\log X_1)/(\log X_2)$  is an ancillary statistic.