

Review:

Stat 522  
1-8-19

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx \quad \leftarrow \text{this is a real number}$$

①

$$E[X|y] = \int_{-\infty}^{\infty} x f(x|y) dx \quad (\text{Defined on 11/27/18})$$

$\nwarrow$  this is a function of  $y$

Defn:  $E[X|Y]$  is a new random variable, and is a function of the r.v.  $Y$

$$\text{Defn: } V[X|Y] = E[(X - E[X|Y])^2 | Y]$$

②

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$$\text{Find } E(E[X|Y])$$

$$= \int_{-\infty}^{\infty} \underbrace{E[X|y]} g(y) dy$$

$$= \int_{-\infty}^{\infty} \left[ \int_{-\infty}^{\infty} x f(x|y) dx \right] g(y) dy$$

$$= \int_{-\infty}^{\infty} \left[ \int_{-\infty}^{\infty} x \frac{f(x,y)}{g(y)} dx \right] g(y) dy$$

(3)

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x, y) dx dy$$

$$= \int_{-\infty}^{\infty} \left[ \int_{-\infty}^{\infty} f(x, y) dy \right] x dx$$

$$= \int_{-\infty}^{\infty} f_x(x) x dx = E[X]$$

$$\therefore E(E[X|Y]) = E[X]$$

"Law of iterated expectations"

(4)

$$V[X] = E[(X - \mu_x)^2]$$

$$= E[(X - E[X|Y] + E[X|Y] - \mu_x)^2]$$

$$= E[(X - E[X|Y])^2] \quad (1)$$

$$+ E[(E[X|Y] - \mu_x)^2] \quad (2)$$

$$+ 2E[(X - E[X|Y])(E[X|Y] - \mu_x)] \quad (3)$$

⑤

$$\begin{aligned}
 \textcircled{3}: & 2 \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x - E[X|y])(E[X|y] - \mu_x) f(x,y) dx dy \\
 &= 2 \int_{-\infty}^{\infty} (E[X|y] - \mu_x) \left[ \underbrace{\int_{-\infty}^{\infty} (x - E[X|y]) f(x,y) dx}_{\int_{-\infty}^{\infty} x f(x,y) dx - E[X|y] \int_{-\infty}^{\infty} f(x,y) dx} \right] dy = 0 \\
 &\quad \underbrace{\int_{-\infty}^{\infty} x f(x,y) dx}_{g(y) E[X|y]} - E[X|y] \underbrace{\int_{-\infty}^{\infty} f(x,y) dx}_{g(y)}
 \end{aligned}$$

⑥

$$\textcircled{2}: E[(E[X|Y] - \mu_x)^2]$$

$$\text{let } W = E[X|Y]$$

$$= E[(W - E[W])^2]$$

$$\begin{aligned}
 \text{Then } E[W] &= E[E[X|Y]] \\
 &= E[X] \\
 &= \mu_x
 \end{aligned}$$

$$= V[W]$$

$$= V[E[X|Y]]$$

$$\textcircled{1}: E[(X - E[X|Y])^2] = E\left(E[(X - E[X|Y])^2 | Y]\right)$$

(using law of iterated expectations)

(7)

$$\textcircled{1}: E[V(X|Y)]$$

$$\therefore V(X) = E[V(X|Y)] + V[E(X|Y)]$$


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Suppose that X and Y are independent r.v.s

Let  $g(X)$  and  $h(Y)$  be 2 new r.v.s

$$\text{Then } E[g(X)h(Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y) \underbrace{f(x,y)}_{f_x(x)f_y(y) \text{ by indep}} dx dy$$

$$= \int_{-\infty}^{\infty} h(y)f_y(y) \left[ \underbrace{\int_{-\infty}^{\infty} g(x)f_x(x) dx}_{E[g(X)]} \right] dy$$

(8)

$$= E[g(X)] E[h(Y)]$$

In particular, if  $X \neq Y$  are ~~not~~ independent,

$$\text{then } E[XY] = E[X]E[Y].$$


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Defn: The covariance of X and Y is

$$E[(X-\mu_X)(Y-\mu_Y)] = \text{Cov}(X,Y) = \sigma_{XY}$$

(9)

Alternate formula:

$$\begin{aligned}
 \sigma_{xy} &= E[(X - \mu_x)(Y - \mu_y)] \\
 &= E[XY - X\mu_y - \mu_x Y + \mu_x \mu_y] \\
 &= E[XY] - \mu_y \underbrace{E[X]}_{\mu_x} - \mu_x \underbrace{E[Y]}_{\mu_y} + \mu_x \mu_y \\
 &= E[XY] - \mu_x \mu_y
 \end{aligned}$$

Note:  $\text{Cov}(X, X) = E[X^2] - \mu_x^2 = V[X]$

(10)

Note: If  $X$  and  $Y$  are independent, then

$$\sigma_{xy} = E[XY] - \mu_x \mu_y = E[X]E[Y] - \mu_x \mu_y = 0$$

Defn: The correlation between  $X$  and  $Y$  is

$$\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{V[X]V[Y]}}$$

$$\rho_{xy} = \frac{\sigma_{xy}}{\sigma_x \sigma_y}$$

(11)

Let  $X$  and  $Y$  be independent r.v.s

Let  $W = X + Y$

Find the moment generating function of  $W$

$$\begin{aligned} M_W(t) &= E[e^{tW}] \\ &= E[e^{t(X+Y)}] = E[e^{tX} e^{tY}] \\ &= E[e^{tX}] E[e^{tY}] = M_X(t) M_Y(t) \end{aligned}$$

(12)

Example: Let  $X$  have a Poisson distribution with parameter  $\mu_1$ .

Let  $Y$  have a Poisson distribution with parameter  $\mu_2$ , independent of  $X$

Let  $W = X + Y$

$$\begin{aligned} M_W(t) &= M_X(t) M_Y(t) = e^{\mu_1(e^t - 1)} e^{\mu_2(e^t - 1)} \\ &= e^{(\mu_1 + \mu_2)(e^t - 1)} \end{aligned}$$

So  $W \sim \text{Poisson}(\mu_1 + \mu_2)$