

Stat 563
4-8-24

The Rao-Blackwell Theorem

Let $X_1, \dots, X_n \sim \text{iid } f(x|\theta)$

①

Let T be a sufficient statistic for θ .

Let W be an unbiased estimator of $\tau(\theta)$.

Let $\phi(T) = E[W|T]$

Then $\phi(T)$ is an unbiased estimator of $\tau(\theta)$

and $V[\phi(T)] \leq V[W] \forall \theta$.

②

$$\begin{aligned}\text{Proof: } E[\phi(T)] &= E[E[W|T]] \\ &= E[W] = \tau(\theta)\end{aligned}$$

$$V[W] = \underbrace{E[V[W|T]]}_{\geq 0} + \underbrace{V[E[W|T]]}_{V[\phi(T)]}$$

$$V[W] \geq V[\phi(T)]$$

Also, $f(\bar{x}|T)$ is free of θ , so
distribution of $W|T$ is free of θ ,
as is $\phi(T) = E(W|T)$.

Example: $X_1, X_2 \sim \text{iid } \frac{1}{\theta} e^{-\frac{x}{\theta}} \quad x > 0$ ③

$$\text{let } W = X_1 \quad E[W] = \theta \quad V[W] = \theta^2$$

Find a sufficient statistic for θ .

$$\begin{aligned} L(\theta) &= \frac{1}{\theta} e^{-\frac{x_1}{\theta}} \frac{1}{\theta} e^{-\frac{x_2}{\theta}} \\ &= \frac{1}{\theta^2} e^{-\frac{1}{\theta}(x_1 + x_2)} \cdot 1 \end{aligned}$$

So $T = X_1 + X_2$ is sufficient for θ .

Find $E[W|T]$ ④

Start with the joint distribution of W, T .

$$W = X_1$$

$$X_1 = w$$

$$T = X_1 + X_2$$

$$X_2 = T - w$$

$$|J| = \begin{vmatrix} 1 & 0 \\ -1 & 1 \end{vmatrix} = 1$$

$$\begin{array}{l} X_1 > 0 \\ X_2 > 0 \\ \hline W > 0 \\ t - w > 0 \\ t > w \end{array}$$

$$\begin{aligned} g(w, t) &= f(x_1, x_2) \cdot 1 = \frac{1}{\theta^2} e^{-\frac{1}{\theta}(x_1 + x_2)} \\ &= \frac{1}{\theta^2} e^{-\frac{t}{\theta}} \quad 0 < w < t \end{aligned}$$

$$g(w|t) = \frac{g(w,t)}{h(t)} = \frac{\frac{1}{\theta^2} e^{-\frac{w}{\theta}}}{\frac{1}{\theta^2} t e^{-\frac{t}{\theta}}} \quad (5)$$

Because $T \sim \text{Gamma}(\alpha=2, \beta=\theta)$

$$= \frac{1}{t}, \quad 0 < w < t$$

That is, $W|T \sim \text{Unif}(0, T)$

$$\text{So } E[W|T] = \frac{0+T}{2} = \frac{T}{2} = \frac{X_1+X_2}{2} = \bar{X}$$

$$\text{Note } E[\bar{X}] = \theta \text{ and } V[\bar{X}] = \frac{\theta^2}{2}$$

A simpler approach might be to (6)

- ① Find a sufficient statistic for θ , called T
- ② Find a function of T that is an unbiased estimator of θ .

Defn: W is a ^(MVUE) minimum variance unbiased

estimator of θ if $E(W)=\theta$ and its variance is less than or equal to the variance of any other unbiased estimator of θ , $V\theta$.

(7)

Theorem: If W is a MVUE of Θ ,
then W is unique.

Proof: let W' also be MVUE of Θ .

$$\text{Let } W^* = \frac{1}{2}(W + W')$$

$$E[W^*] = \frac{1}{2}[E(W) + E(W')] = \Theta$$

$$\begin{aligned} V[W^*] &= \frac{1}{4}[\underline{V[W] + V[W']} + 2\text{Cov}(W, W')] \\ &= \frac{1}{4}[2V[W] + 2\text{Cov}(W, W')] \end{aligned}$$

$$V[W^*] = \frac{1}{2}V[W] + \frac{1}{2}\text{Cov}(W, W')$$

(8)

$$\text{Recall } \rho^2 = \frac{\text{Cov}^2(W, W')}{V[W]V[W']} \leq 1$$

$$\text{So } \text{Cov}^2(W, W') \leq (V[W])^2$$

$$|\text{Cov}(W, W')| \leq V[W] \quad \star$$

$$\text{Now } V[W^*] \leq \frac{1}{2}V[W] + \frac{1}{2}V[W']$$

$$V[W^*] \leq V[W]$$

It, for any θ , $V(W^*) < V(W)$, ⑨

our assumption that W was MVUE
would be violated.

$$\therefore V(W^*) = V(W) + \theta$$

Equivalently, $\rho = \text{Cor}(W, W') = 1$

(Cauchy-Schwarz says $W' = aW + b$)

By *, $\text{Cov}(W, W') = V(W)$

$\text{Cov}(W, W') = \text{Cov}(W, aW + b)$ ⑩

$$= a \text{Cov}(W, W) + \text{Cov}(W, b)$$

$$= aV(W) + 0$$

$$\therefore a = 1$$

So $W' = W + b$

$$E(W') = E(W) + b$$

$$\theta = \Theta + b \quad \therefore b = \theta$$

$$\therefore W = W'$$

(11)

The Lehmann - Scheffé Theorem

Let T be a complete sufficient statistic for θ . If there exists W , a function of T , that is an unbiased estimator of $\tau(\theta)$, then W is the unique MVUE of $\tau(\theta)$.

Proof: Assume W exists.

let W' be another unbiased estimator of $\tau(\theta)$.

(12)

By the Rao-Blackwell Theorem,

$E[W'|T]$ will be an unbiased estimator of $\tau(\theta)$ and its variance will be $\leq V[W']$.

Also note that $E[W'|T]$ is a function of T .

Let $\varphi(T)$ and $\psi(T)$ be functions of T that are both unbiased estimators of $\tau(\theta)$.

$$\begin{aligned} E[\Psi(T) - \bar{\Psi}(T)] &= E[\Psi(T)] - E[\bar{\Psi}(T)] \\ &= \gamma(\theta) - \bar{\gamma}(\theta) = 0 \quad \forall \theta \end{aligned}$$

By completeness, this implies

$$\Psi(T) - \bar{\Psi}(T) = 0$$

So W and $E[W|T]$ must be equal.

$$\text{Then } V(W) = V[E(W|T)] \leq V(W)$$

$\therefore W$ is MVUE and is unique.