Time series as a stochastic process must satisfy the property of stationarity and ergodicity

Stationarity

 A time series process {z_i} is (strongly) stationary if the joint probability distribution of any set of k observations in the sequence {z_i, z_{i+1},..., z_{i+k-1}} is the same regardless of the origin, i, in the time scale.

Stationarity

 A time series process {z_i} is covariance stationary or weakly stationary if $E(z_i) = \mu$ is finite and is the same for all i; and if the covariance between any two observations, $Cov(z_i, z_{i-k}) = \gamma_k$, is a finite function only of model parameters and their distance apart in time k, but not the absolute location of either observation on the time scale.

•
$$\gamma_k = \gamma_{-k}$$

Ergodicity

 A stationary time series process {z_i} is ergodic if for any two bunded functions that map vectors in the a and b dimensional real vector spaces to real scalars, f: R^a→ R and g: R^b→ R,

$$\begin{split} \lim_{k \to \infty} \left| E \Big[f \Big(z_i, z_{i+1}, \dots, z_{i+a-1} \Big) g \Big(z_{i+k}, z_{i+k+1}, \dots, z_{i+k+b-1} \Big) \Big] \right| \\ = \left| E \Big[f \Big(z_i, z_{i+1}, \dots, z_{i+a-1} \Big) \Big] \right| \left| E \Big[g \Big(z_{i+k}, z_{i+k+1}, \dots, z_{i+k+b-1} \Big) \Big] \right| \end{split}$$

Ergodic Theorem

 If {z_i} is a time series process that is stationary and ergodic and E[|z_i|] is a finite constant, and if

$$\overline{z}_n = \frac{1}{n} \sum_{i=1}^n z_i, \text{ then } \overline{z}_n \to_{as} \mu, \text{ where } \mu = E(z_i)$$

Ergodicity of Functions

- If {z_i} is a time series process that is stationary and ergodic and if y_i = f(z_i) is a measurable function in the probability space that defines z_i, then {y_i} is also stationary and ergodic.
- Let {Z_i} define a Kx1 vector valued stochastic process—each element of the vector {z_i} is an ergodic and stationary series, and the characteristics of ergodicity and stationarity apply to the joint distribution of the elements of {Z_i}. Then, the ergodic theorem applies to functions of {Z_i}.

Martingale Sequence

- A vector sequence {Z_i} is a martingale sequence if E(Z_i | Z_{i-1}, Z_{i-2},...) = Z_i.
- A vector sequence {Z_i} is a martingale difference sequence (m.d.s) if E(Z_i | Z_{i-1}, Z_{i-2},...) = 0.

Martingale Difference CLT

• If $\{Z_i\}$ is a vector valued stationary and ergodic martingale difference sequence, with $E(Z_iZ_i') = \Sigma$, where Σ is a finite positive definite matrix, and if

$$\overline{Z}_n = \frac{1}{n} \sum_{i=1}^n Z_i, \text{ then } \sqrt{n} \overline{Z}_n \to_d N(0, \Sigma)$$

Gordin's CLT

• If {Z_i} is stationary and ergodic and if the following three conditions are met, then

$$\sqrt{n}\overline{Z}_{n} \to_{d} N(0, \Gamma^{*})$$
where $\Gamma^{*} = \lim_{n \to \infty} Var\left(\sqrt{n}\overline{Z}_{n}\right)$

- 1. Asymptotic uncorrelatedness: $E(Z_i | Z_{i-k}, Z_{i-k-1}, ...)$ converges in mean squares to zero as $k \rightarrow \infty$.

Gordin's CLT

2. Summability of autocovariances:
 With dependent observations, the following covariance matrix is finite.

$$\lim_{n \to \infty} Var\left(\sqrt{n}\overline{Z}_n\right) = \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} Cov(Z_i, Z_j) = \sum_{k=-\infty}^{\infty} \Gamma_k = \Gamma^*$$

In particular, $E(Z_iZ_j') = \Gamma_0$, a finite matrix.

Gordin's CLT

 Asymptotic negligibility of innovations:
 The information eventually becomes negligible as it fades far back in time from the current observation.

Let
$$r_{ik} = E(Z_i | Z_{i-k}, Z_{i-k-1}, ...) - E(Z_i | Z_{i-k-1}, Z_{i-k-2}, ...),$$

then $Z_i = \sum_{k=0}^{\infty} r_{ik}$ and $\sum_{k=0}^{\infty} \sqrt{E(r_{ik} r_{ik})}$ is finite

- Linear Model: $y_i = \mathbf{x}_i' \mathbf{\beta} + \varepsilon_i$ (i=1,2,...,n)
- Least Squares Estimation

$$-b = (X'X)^{-1}X'y = \beta + (X'X/n)^{-1}(X'\epsilon/n)$$

$$- E(b) = \beta + (X'X/n)^{-1} E(X'\varepsilon/n) = \beta$$

- Var(b) = $(X'X/n)^{-1} E[(X'\epsilon/n)(X'\epsilon/n)'](X'X/n)^{-1}$ = $(X'X)^{-1}(X'\Omega X)(X'X)^{-1}$ where $\Omega = E(\epsilon\epsilon')$

- We assume the stochastic process generating $\{\mathbf{x}_i\}$ is stationary and ergodic, then by ergodic theorem, $\lim_{n\to\infty} \sum_i \mathbf{x}_i \mathbf{x}_i'/n = E(\mathbf{x}_i \mathbf{x}_i')$ is finite
- If ε_i is not serially correlated, then g_i = x_iε_i is a m.d.s, so
 - $-\lim_{n\to\infty}\sum_{i}\mathbf{x}_{i}\varepsilon_{i}/n = E(\mathbf{x}_{i}\varepsilon_{i}) = E(\mathbf{g}_{i})=\mathbf{0}$
 - $-\lim_{n\to\infty}\sum_{i}\mathbf{x}_{i}\mathbf{x}_{j}'\epsilon_{i}\epsilon_{j}/n = E(\mathbf{g}_{i}\mathbf{g}_{j}') = Var(\mathbf{g}_{i}) \text{ is finite}$
- Therefore, b is consistent!

- If {x_i,ε_i} are jointly stationary and ergodic, then by ergodic theorems, consistency of the parameter estimates can be established.
- If ϵ_i is not serially correlated, then by Martingale Difference CLT,

 $\overline{\mathbf{g}}_n = \frac{1}{n} \sum_{i=1}^n \mathbf{g}_i, \text{ then } \sqrt{n} \overline{\mathbf{g}}_n \rightarrow_d N(\mathbf{0}, \boldsymbol{\Sigma}) \text{ where } \boldsymbol{\Sigma} = \mathbf{X}' \boldsymbol{\Omega} \mathbf{X}$

 If ε_i is serially correlated and dependence in x_i, then by Godin's CLT,

$$\sqrt{n}\overline{\mathbf{g}}_n \rightarrow_d N(0,\Gamma^*), where \Gamma^* = \lim_{n \to \infty} Var\left(\sqrt{n}\overline{\mathbf{g}}_n\right)$$